1980 Census of Population and Housing DOT Journey-to-Work (SMSA) 1975-1976 Data Base Dictionary.

D1-D80-WORK-14-TECHDD

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U.S. Department of Commerce Economic and Statistics Administration U.S. Census Bureau



DOT JOURNEY TO WORK (SMSA) 1975-76

DATA BASE DICTIONARY

DATA BASE DICTIONARY

TEXT SECTION

LOGICAL RECORD LENGTH = 408

FILE NAME : DOT JOURNEY TO WORK [SNSA] 1975-76

INTRODUCTION:

THE TRAVEL TO WORK DATA CONTAINED IN THIS FILE ARE FROM THE TRAVELTO-WORK SUPPLEMENT TO THE 1975-76 ANNUAL HOUSING SURVEY [AHS] SMSA
SAMPLE. SPONSORED BY THE US DEPARTMENT OF TRANSPORTATION. THE 1975-76
AHS SMSA SAMPLE WAS ENUMERATED BY DIRECT PERSONAL INTERVIEW, DURING
THE PERIOD APRIL 1975 THROUGH MARCH 1976, BY THE BUREAU OF THE CENSUS URBAN DEVELOPMENT. ACTING AS THE COLLECTION AGENT FOR THE US DEPARTMENT OF HOUSING AND

UNITS LOCATED IN EACH OF FOUR LARGE SMSA'S AND 5.000 IN THE REMAINING 17 METROPOLITAN AREAS. THE SAMPLE WAS SELECTED FROM HOUSING CHAIRS IN THE 1970 CENSUS AND UPDATED BY A SAMPLE OF ADDRESSES FROM BUILDING PERMITS TO INCLUDE UNITS CONSTRUCTED SINCE 1970. HOUSING UNITS OCCUPIED ENTIRELY BY PERSONS WITH USUAL RESIDENCE ELSEWHERE [URE] ARE VACANT UNITS. THE CHARACTERISTICS OF THE TEMPORARY RESIDENTS DO NOT APPEAR ON THE FILE. VACANT OR URE MOBILE HOMES WITH NO PERMANENT ROOM ADDED. VACANT OR URE PERMANENT HOUSING UNITS IN TRANSIENT HOTELS OR MOTELS. AND VACANT OR URE UNSPECIFIED HOUSING UNITS ARE TREATED AS THE 1975-76 AMS SMSA SAMPLE CONSISTS OF APPROXIMATELY 15.000 HOUSING NONINTERVIEWS.

THE TRAFEL-TO-WORK SUPPLEMENT IS DESIGNED TO PROVIDE CURRENT INFORMATION FOR SELECTED METROPOLITAN AREAS CONCERNING SUCH SUBJECTS AS PLACE OF NORK. MEANS OF TRANSPORTATION TO WORK, CARPOOLING AND CARPOOL OCCUPANCY. TRAVEL TIME TO WORK, DISTANCE TO WORK, TIME OF DEPARTURE FOR WORK, AND OTHER TRANSPORTATION-RELATED SUBJECTS. THE SERIES OF SUPPLEMENTARY QUESTIONS ON TRAVEL TO WORK WERE ASKED OF ALL HOUSEHOLD MEMBERS 14 YEARS AND OVER WHO HAD A JOB DURING THE WEEK PRIOR TO INTERVIEW. THIS PUBLIC USE COMPUTER TAPE PRESENTS THE BASIC RECORDS OF EACH OF THE APPROXIMATELY 267.000 WORKERS SAMPLED IN THE 1975-76 TRAVEL-TO-WORK SUPPLEMENT. THE TILE CONTAINS TRANSPORTATION DATA FROM THE SUPPLEMENT, AND SCLECTED DEMOGRAPHIC AND HOUSEHOLD CHARACTERISTICS WHICH ARE ALSO AVAILABLE FROM THE AHS PUBLIC USE FILES. A LIST OF THE DATA ITEMS AND THEIR UNIVERSES IS AVAILABLE WITH THIS DOCUMENTATION.

GEOGRAPHIC IDENTIFICATION PROVIDED ON THE FILES IS IN ACCORDANCE WITH THE BUREAU OF THE CENSUS' CONFIDENTIALITY RESTRICTIONS. WHICH BY LAW [TITLE 13. US CODE] PROHIBIT THE DISCLOSURE OF ANY DATA WHICH MIGHT SPECIFICALLY IDENTIFY ANY INDIVIDUAL PERSON. HOUSEHOLD. OR HOUSING UNIT. AS THESE DATA ARE BASIC WORKER RECORDS. THE ADVANTAGES DERIVED FROM THESE TAPE FILES INCLUDE THE CAPABILITY OF OBTAINING A MORE DETAILED CLASSIFICATION WITHIN EACH OF THE ITEMS SHOWN. AS WELL AS THE FLEXIBILITY OF MANIPULATING VARIABLES TO FIT A USER'S SPECIFIC

SMSA IDENTIFICATION:

TEXT SECTION

THE FOLLOWING SMSA'S AND SMSA COMPONENTS ARE CONTAINED ON THE 1975-76 SMSA FILES:

NEW ORLEANS. LA	MILWAUKEE. WI MILWAUKEE BALANCE OF SMSA	MIAMI. FL	MADISON. WI	KANSAS CITY, MO-KS	HARTFORD.CN	COLUMBUS.OH	COLORADO SPRINGS, CO	CINCINNATI OH-KY-IN	SAN FRANCISCO-OAKLAND. CA ALAMEDA SAN FRANCISCO SAN MATEO BALANCE OF SMSA	CHICAGO.IL COOK DUPAGE KANE LAKE BALANCE OF SMSA	ATLANTA.GA DEKALB.FULTON BALANCE OF SMSA	NEW JERSEY CAMDEN BALANCE OF SMSA		COUNTY PHILADELPHIA.PA-NJ PENNSYLVANIA BUCKS
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TEXT SECTION

	SPRINGFIELD-CHICOPEE- HOLYOKE, MA-CN	SAN DIEGO.CA	SAN BERNARDINO-RIVERSIDE- ONTARIO. CA	SAN ANTONIO. TX	ROCHESTER. NY	PORTLAND. OR-WA	PATERSON-CLIFTON-PASSAIC. NJ	NEWPORT NEWS-HAMPTON, VA	ORLEANS BALANCE OF SMSA
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SAMPLING VARIABILITY ASSOCIATED WITH THESE DATA IS DISCUSSED IN APPENDIX B ATTACHED TO THIS TECHNICAL DOCUMENTATION.

USE OF THIS : DATA BASE DICTIONARY:

THIS COMPUTER GENERATED REPORT DOCUMENTS THE DATA CONTENTS AND RECORD LAYOUT FOR THE TRAVEL-TO-WORK SUPPLEMENT TO THE 1975-76 SMSA AHS. THIS REPORT IS ORGANIZED INTO 2 PARTS, ONE PART CONTAINING A GENERAL TEXTUAL DESCRIPTION OF THE FILE. AND ONE PART CONTAINING A DESCRIPTION OF EACH DATA ITEM IN THE FILE. THE DATA ITEM DESCRIPTION CONTAINS SEVERAL PIECES OF INFORMATION FOR EACH DATA ITEM AS FOLLOWS:

AN ARBITRARILY ASSIGNED 12 CHARACTER LABEL. IT MAY BE A MNEMONIC OR A NUMERIC VALUE INDICATING THE LOCATION OF THE VARIABLE ON THE SURVEY QUESTIONNAIRE. FOR EXAMPLE. HOOG IS A HOUSING VARIABLE DERIVED FROM THE QUESTIONNAIRE INDICATED BY THE NUMBER 6 ENCIRCLED ON THE QUESTIONNAIRE

NAME:

DESCRIPTIVE

LABEL: A 40 CHARACTER LABEL DESCRIBING THE DATA ITEM

LENGTH: THE SIZE IN CHARACTERS, OF THE DATA ITEM

BEGIN: THE LOCATION IN THE DATA RECORD OF THE FIRST CHARACTER OF THE DATA ITEM

MAXIMUM VALUE: THE HIGHEST VALUE THIS DATA ITEM MAY CONTAIN

MINIMUM VALUE: THE LOWEST VALUE THIS DATA ITEM MAY CONTAIN

TEXT SECTION

NOTE: MAXIMUM AND MINIMUM VALUES ARE DISPLAYED ONLY FOR CODE TYPE ITEMS [SUCH AS STATE CODE]

DATA CATEGORY: THIS FIELD INDICATES WHETHER THE DATA ITEM IS A NUMERIC ITEM WHICH CAN BE MEANINGFULLY PROCESSED ALGEBRAICALLY OR A CODE ITEM FROM WHICH IT WOULD NOT BE MEANINGFUL TO SUMMARIZE OR DERIVE PERCENTAGES

DECIMAL PLACES: THIS FIELD INDICATES HOW MANY, IF ANY, IMPLIED DECIMAL PLACES THE DATA ITEM CONTAINS

FILE STRUCTURE:

THE FILE SET INCLUDES SEPARATE FILES FOR EACH OF THE 21 SMSAS CONTAINED IN THE 1975-76 SURVEY. WITHIN EACH SMSA THE BASIC UNIT IS A WORKER RECORD THAT IS. THERE IS A SIPARATE DATA RECORD FOR EACH WORKER IN THE SURVEY. IN ADDITION TO THE SPICIFIC WORKER'S TRANSPORTATION CHARACTERISTICS, EACH RECORD ALSO CONTAINS THE FULL RANGE OF CHARACTERISTICS ASSOCIATED WITH HIS OR HER OWN HOUSEHOLD AND HOUSEHOLD THE DATA RECORDS ARE ARRANGED IN HOUSEHOLD CONTROL NUMBER ORDER WITH UNIQUE IDENTIFICATION BEING PROVIDED BY THE COMBINATION OF CONTROL NUMBER AND WORKER LINE NUMBER. IN CASES WHERE THERE ARE NO WORKERS IN THE HOUSEHOLD. THERE IS A RECORD PROVIDED WHICH CONTAINS DEMOGRAPHIC AND HOUSEHOLD CHARACTERISTICS BUT THE WORKER WEIGHT AND LINE NUMBER OF THE WORKER FIELDS ARE ZEROBUT THE TRANSPORTATION DATA ITEM FIELDS ARE FILLED WITH 9'S.

RECORDING MODE:

THESE FILES WILL NORMALLY BE RECORDED ON TAPE WITH :IBM: STANDARD LABELS OPTIONALLY THE FILES CAN BE SUPPLIED ON UNLABELLED TAPES WITH A SINGLE TAPE MARK SEPARATING THE FILES.

CHARACTER SET:

A LIMITED CHARACTER SET CONSISTING OF ALPHABETIC AND NUMERIC CHARACTERS ONLY IS USED IN THE DATA FILES. SEVEN TRACK TAPES WILL BE RECORDED USING THE BCD CHARACTER SET AND 9 TRACK TAPES WILL BE RECORDED USING THE EBCDIC CHARACTER SET.

FILE SIZE AND COST:

DATA FOR EACH METROPOLITAN AREA [SMSA] ARE CONTAINED IN A SERIES DF 408 CHARACTER RECORDS, OME RECORD FOR EACH OF THE WORKERS IN THE SURVEY. DATA FOR A SINGLE SMSA CAN BE PURCHASED SEPARATELY OR ALL 21 FILES CAN BE OBTAINED ON 2 OR 4 REELS OF TAPE DEPENDING THE DESIRED DENSITY. THE COST OF EACH REEL IS \$80.

THESE FILES CAN BE ORDERED THROUGH:

TEXT SECTION

BUREAU OF THE CENSUS WASHINGTON: D.C. 20233 PHONE [301] 763-2400 DATA USERS SERVICES DIVISION CUSTOMER SERVICES BRANCH

THE DATA ARE ALSO AVAILABLE RECORDED WITH SMALLER BLOCKSIZES FOR THOSE PURCHASERS WHO REQUIRE SMALLER BLOCKSIZES. ADDITIONAL REELS OF TAPE WAY BE REQUIRED WHEN COPIES ARE MADE USING SMALLER BLOCKSIZES. WHEN ORDERING PLEASE SPECIFY DENSITY AND TRACK DESIRED. PAYMENT MUST BE INCLUDED WITH THE ORDER.

FURTHER INFORMATION:

FOR PUBLISHED DATA FROM THIS SURVEY SEE:

SELECTED CHARACTERISTICS OF TRAVEL TO WORK IN 21 METROPOLITAN AREAS: 1975. SERIES P-23. NUMBER 68.

FOR FURTHER INFORMATION AND COMMENTS PLEASE CONTACT :

MR PHILIP FULTON, CHIEF, JOURNEY-TO-WORK STATISTICS STAFF POPULATION DIVISION

PHILLIP SALOPEK. JOURNEY-TO-WORK STATISTICS STAFF POPULATION DIVISION

MR. LARRY CARBAUGH DATA USER SERVICES DIVISION

NOTES:

FIELDS CONTAINING INCOME VALUES [H241, H281, ETC] ARE TOO LARGE IN TERMS OF THE NUMBER OF CHARACTERS TO APPLY THE FOUR CHARACTER LABEL ALLOWED FOR VALUE LABELS IN THE DICTIONARY. THE DATA ARE THUS REPRESENTED AS FOLLOWS..

00000 00001 NO INCOME

49999 \$49.999
50000 \$50.000 OR MORE
50000 \$50.000 OR MORE
99999 OUT OF UNIVERSE
FIELDS H252. H254 [SELF EMPLOYMENT INCOME] CAN ALSO HAVE NEGATIVE
VALUES SO THAT .10000 REPRESENTS AN INCOME LOSS OF \$10.000.

SPECIAL TABULATION ZONES:

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XT SECTION

THE BUREAU OF THE CENSUS HAS DEFINED SPECIAL TABULATION ZONES FOR THE PURPOSE OF REPORTING THE RESULTS OF THE DEPARTMENT OF TRANSPORTATION TRAVEL-TO-WORK SUPPLEMENT. EACH SMSA HAS BEEN DIVIDED ALONG TRACT BOUNDARIES INTO AT LEAST 2 SPECIAL TABULATION ZONES, LABELED A AND R . SPECIAL TABULATION ZONE A (TZA) WAS DELINEATED ACCORDING TO THE CRITERIA OUTLINED BELOW. SPECIAL TABULATION ZONE R (TZR) CONSISTS OF THE REMAINDER OF THE TRACTS IN THE SMSA, THAT IS. THOSE THAT DO NOT MEET THE CRITERIA. IN ALL CASES THE CENSUS TRACT BOUNDARIES AND INCORPORATED PLACE POPULATIONS USED IN DEFINING THE SPECIAL TABULATION ZONES WERE THOSE OBSERVED FOR OR RESULTING FROM THE 1970 CENSUS.

IN ORDER TO ACHIEVE A NATIONALLY UNIFORM SET OF AREAS FOR USE IN THE TRAVEL-TO-WORK SUPPLEMENT. THE OUTER BOUNDARY OF TZA (WITH A FEW EXCEPTIONS) ENCOMPASSES THOSE TRACTS WHICH IN 1970 HAD 40% OR MORE OF THEIR POPULATION LIVING INSIDE THE URBANIZED AREA (UA) AS DEFINED BY THE BUREAU OF THE CENSUS. THE FOLLOWING EXCEPTIONS WERE MADE TO THE GENERAL RULE IN ORDER TO MAKE THE TZA MORE COMPACT AND TO REDUCE THE JAGGEDNESS OF THE BOUNDARY.

- TRACTS WITH NO POPULATION LIVING INSIDE THE 1970 URBANIZED AREA WERE INCLUDED WITHIN THE TZA IF THEY WERE COMPLETELY SURROUNDED BY QUALIFYING TRACTS.
- 2. TRACTS WITH LESS THAN 20% OF THEIR TOTAL POPULATION LIVING WITHIN THE 1970 UN WERE INCLUDED WITHIN THE TZA IF A) THEY WERE COMPLETELY SURROUNDED BY QUALIFYING TRACTS: OR B) THEY INCLUDED 500 OR MORE INHABITANTS OF THE SMSA CENTRAL CITY AS BOUNDED JANUARY 1, 1970.
- 3. TRACTS WITH 20 TO 40 PERCENT OF THEIR TOTAL POPULATION LIVING WITHIN THE 1970 UA WERE INCLUDED WITHIN TZA IF A] THEY WERF COMPLETELY SURRGUNDED BY QUALIFYING TRACTS B] THEY INCLUDED 5:00 OR MORE INHABITANTS OF AN SMSA CENTRAL CITY [AS BOUNDED JANUARY 1 1970]. OR C] THE INCLUSION OF THE TRACT SERVED TO CLOSE AN INDENTATION AND PROVIDE A SMOOTHER TZA BOUNDARY.
- 4. IN CASES WHERE THE TRACT HAS FAILED TO QUALIFY FOR INCLUSION WITHIN TZA UNDER THE CRITERIA OUTLINED ABOVE. BUT THERE IS A PART OF THE TRACT WHICH IS INSIDE THE CITY LIMITS. THE TZA BOUNDARY DOES NOT FOLLOW TRACT LINES. IN THESE INSTANCES THE TZA BOUNDARY FOLLOWS THE CITY LIMITS. SO THAT THE CENTRAL CITY PART OF THE TRACT IS INCLUDED IN TZA. AND THE REMAINDER OF THE TRACT IS INCLUDED IN TZA.

IN A FEW INSTANCES THE SMSA HAS BEEN DIVIDED INTO MORE THAN TWO SPECIAL TABULATION ZONES. WHEN THIS OCCURS, THE CRITERIA OUTLINED ABOVE HAVE BEEN EMPLOYED TO DEFINE AS MANY AS FOUR AREAS [TZA, TZB, TZC, OR TZD], WITH THE REMAINDER OF THE TRACTS IN THE SMSA AGAIN

TEXT SECTION

COMPRISING TZR. QUESTIONS CONCERNING THE DELINEATION OF THE SPECIAL TABULATION ZONES, INCLUDING REQUESTS FOR MAPS INDICATING THE EXACT BOUNDARIES USED, MAY BE REFERRED TO THE JOURNEY TO WORK STATISTICS STAFF, POPULATION DIVISION, BUREAU OF THE CENSUS.

LINE NUMBERS:

SELECTED DATA [AGE. SIX. RACE. MARITAL STATUS. RELATIONSHIP TO HEAD AND WAGE/SALARY INCOME] ARE PRESENTED ON EACH RECORD FOR EACH PERSON IN A HOUSING UNIT. EACH PERSON [UP TO 15 PEOPLE] IS ASSIGNED A LINE NUMBER. THE DATA FIELDS ARE IDENTIFIED WITH THIS NUMBER [E.G. LNNO2 IS THE SECOND PERSON IN THE UNIT. AGEOZ CONTAINS THE AGE OF THIS PERSON. SEXOZ CONTAINS THE SEX. ETC.]. H240, H242. H244, H246, H248, AND H250 IDENTIFY FAMILY MEMBERS 14 YEARS OLD AND COVER WITH H241. H243. H245, H247. H249 AND H251 A YEARS OLD AND ARE NOT AVAILABLE FOR PERSONS IN THE HOUSEHOLD WHO ARE NOT RELATED TO THE HOUSEHOLD HEAD]. IF THERE ARE MORE THAN SIX PERSONS 14 YEARS OLD AND OVER IN THE FAMILY FIELD H250 WILL CONTAIN A 99 AND H251 WILL BE THE AGGREGATE INCOME OF THE REMAINING PERSONS. THAT IS, PERSON 6, PERSON 7, ETC.

LINE NUMBER OF WORKER: THIS LINE NUMBER, IN CONJUNCTION WITH THE CONTROL NUMBER, UNIQUELY IDENTIFIES THE DATA RECORD AND A PERSON 14 OR OLDER WHO HAD A JOB AND WHOSE TRANSPORTATION CHARACTERISTICS ARE CONZAINED ON THE RECORD. IN ORDER TO PICK UP OTHER CHARACTERISTICS FOR THIS WORKER, THE USER SHOULD MATCH THE WORKER LINE NUMBER TO THE PERSON LINE NUMBER ELSEWHERE ON THE RECORD. IMMEDIATELY FOLLOWING THE APPROPRIATE PERSON LINE NUMBER WILL BE THE RELATIONSHIP, AGE, MARITAL STATUS, RACE, AND SEX OF THE WORKER, THE PERSON LINE NUMBER CAN ALSO BE ASSOCIATED WITH A WAGE/SALARY INCOME FIELD.

LINE NUMBER OF THE DOT RESPONDENT: THIS IS THE LINE NUMBER OF THE PERSON IN THE HOUSEHOLD WHO PROVIDED THE TRAVEL-TO-WORK INFORMATION CONTAINED IN THE RECORD. THIS LINE NUMBER HAS NOT BEEN EDITED FOR CONSISTENCY OR ACCURACY.

RESPONDENT LINE NUMBER: THIS IS THE LINE NUMBER OF THE PERSON IN THE HOUSEHOLD WHO PROVIDED THE HOUSEHOLD AND DEMOGRAPHIC CHARACTERISTICS CONTAINED ON THE RECORD.

WRK393	WRK392		WRK390	RESLNNO	MKLNNO	HHWGT	CNTNO	NAME	
NUMBER OF PERSONS IN CARPOOL O1 ONE PERSON O2 TWO PERSONS O3 THREE PERSONS O4 FOUR PERSONS O5 FIVE PERSONS O6 SIX PERSONS O7 SEVEN PERSONS	WORKER ALSO USES CAR † YES 2 NO 9 OUT OF UNIVERSE	204 AUTO, RIDES WITH SOMEONE ELSE 005 WALKS ONLY 006 WORKS AT HOME 007 RAILROAD 008 SUBWAY OR ELEVATED 009 BUS OR STREETCAR 010 TAXICAB 011 MOTORCYCLE 012 OTHER MEANS 013 BICYCLE 999 OUT OF UNIVERSE	MEANS OF TRUCK. TRUCK. TRUCK. TRUCK. AUTO. AUTO.	LINE NUMBER OF DOT RESPONDENT	ZERO INDICATES NO WORKERS IN HOUSEHOLD LINE NUMBER OF WORKER OO NO WORKER IN HOUSEHOLD	ZERO FOR AL	CONTROL NUMBER SAME AS CONTROL NUMBER ON 1975 AHS SMSA FILLS	DATA SECTION DESCRIPTIVE LABELS	DATA BASE DICTIONARY
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PRIOR MODE OF TRANSPORTATION TO WORK TRUCK. DRIVES ALONE TRUCK, SHARES DRIVING TRUCK, DRIVES OTHERS TRUCK, RIDES WITH SOMEONE ELSE AUTO, DRIVES ALONE	SPEED OF JOURNEY TO WORK [RECODE] O MPH [DISTANCE LESS THAN 1 MILE] 97 MPH OR MORE OUT OF UNIVERSE	MODE CHANGE DURING LAST YEAR YES NO OUT OF UNIVERSE	DISTANCE TO WORK [ONE WAY] LESS THAN 1 MILE 201 MILES OR MORE OUT OF UNIVERSE	TIME TAKEN TO GET TO WORK [ONE WAY] ONE MINUTE 199 MINUTES OR NORE OUT OF UNIVERSE	TIME LEAVING FOR WORK [HOUR AND MINUTE] HOUR AND MINUTE OUT OF UNIVERSE	TIME LEAVING FOR WORK [AM OR PM] AN PN OUT OF UNIVERSE	REPORTS IN TO SAME LOCATION EACH DAY YES NO OUT OF UNIVERSE	WORKS AT SAME LOCATION EACH DAY YES NO OUT OF UNIVERSE	EIGHT PERSONS NINE PERSONS TEN OR MORE PERSONS OUT OF UNIVERSE	DATA SECTION DESCRIPTIVE LABELS	DATA BASE DICTIONARY
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WORKPLACE CENTRAL CITY FLAG 1 CENTRAL CITY 2 OAKLAND. CALIFORNIA 3 BALANCE OF SMSA 6 NO FIXED PLACE OF WORK 7 OUTSIDE SMSA OF RESIDENCE 8 PLACE OF WORK NOT REPORTED	WORKPLACE COUNTY CODE [SEE ATTACHED LIST FOR COUNTY 996 NO FIXED PLACE OF WORK 998 PLACE OF WORK NOT REPORTED 999 OUT OF UNIVERSE	WORKPLACE STATE CODE [SEE ATTACHED LIST FOR STATE 96 NO FIXED PLACE OF WORK 98 PLACE OF WORK NOT REPORTED 99 OUT OF UNIVERSE	WORKPLACE SMSA CODE [SEE ATTACHED LIST FOR SMSA CODES] 9996 NO FIXED PLACE OF WORK 9997 CUTSIDE ANY SMSA 9998 PLACE OF WORK NOT REPORTED 9999 CUT OF UNIVERSE	SATISFACTION WITH MODE OR MUCH MORE SATISFIED MORE SATISFIED ABOUT THE SAME SATISFA LESS SATISFIED MUCH LESS SATISFIED DONT KITOW OUT OF UNIVERSE	202 AUTO. SHARES DRIVING 203 AUTO. DRIVES OTHERS 204 AUTO. RIDES WITH SOMEONE 005 WALKS OF LY 006 WORKS AT HOME 007 RAILROAD 008 SUBWAY OR ELEVATED 019 BUS OR STREETCAR 010 TAXICAB 011 MOTORCYCLE 012 OTHER MEANS 013 BICYCLE 999 OUT OF UNIVERSE	DATA SECTION DESCRIPTIVE LABELS	DATA BASE DICTIONARY
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[SEE ATTACHED LIST FOR COUNTY CODES] NOT APPLICABLE	ECTRENCE COUNTY CODE	COLORADO CONNECTICUT FLORIDA GEORGIA ILLINOIS KANSAS LOUISIANA MASSACHUSETTS MISSOURI NEW JERSEY NEW YORK OHIO PENNSYLVANIA TEXAS VIRGINIIA WISCONSIN KENTUCKY-INDIANA NOT APPLICABLE	RESIDENCE STATE CODE CALIFORNIA	RESIDENCE CENTRAL CITY FLAG INSIDE CENTRAL CITY OAKLAND. CALIFORNIA OUTSIDE CENTRAL CITY NOT APPLICABLE	COLUMBUS. OHIO HARTFORD KANSAS CITY MADISON MIAMI MILWAUKFE MILWAUKFE NEW ORLEANS NEWPORT NEWS-HAMPTON PATERSON-CLIFTON-PASSAIC PHILADELPHIA PORTLAND. OREGON ROCHESTER. NEW YORK SAN ANTONIO SAN BERNARDINO-RIVERSIDE-ONTARIO SAN DIEGO SAN FRANCISCO-OAKLAND SPRINGFIELD-CHICOPEE-HOLYOKE	PTIVE	DATA SECTION	DATA BASE DICTIONARY
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8 008	TYPE OF LIVING QUARTERS OI HOUSE, APT, FLAT O2 UNIT IN NONTRANSIENT HOTEL, ETC O3 UNIT IN TRANSIENT HOTEL, ETC O4 UNIT IN FOOMING HOUSE O5 MOBILE HOME/TRAILER, NO PERM. ROOM O6 MOBILE HOME/TRAILER WITH ROOM ADDED O7 UNIT NOT SPECIFIED ABOVE	2 103 00999 00000	CODE 0
H012	USE OF TELEPHONE 1 YES 2 NO	1 105 00099 00000	CODE 0
H023	TYPE OF STRUCTURE 01 MOBILE HOME, TRAILER 02 ONE UNIT, DETACHED 03 ONE UNIT, ATTACHED 04 2 UNITS 05 3 OR 4 UNITS 06 5 TO 9 UNITS 07 10 TO 19 UNITS 08 20 TO 49 UNITS 09 50 OR MORE UNITS	2 106 00999 00000	CODE 0
H140	PUBLIC HOUSING PROJECT 1 YES HOUSING PROJECT 2 NO HOUSING PROJECT 8 NOT ANSWERED 9 OUT OF UNIVERSE	1 108 00099 00000	CODE
H141	GOVERNMENT SUBSIDIES OF RENT 1 YES SUBSIDIED 2 NO GOVT SUBSIDIES 8 NOT ANSWERED 9 OUT OF UNIVERSE	1 109 00099 00000	CODE 0
H043	GARAGE OR CARPORT AVAILABLE 1 YES 2 NO 8 NOT ANSWERED 9 OUT OF UNIVERSE	1 110 00099 00000	CODE D
H156	PARKING FACILITIES AVAILABLE 1 YES. PARKING AVAILABLE 2 NO PARKING AVAILABLE 8 NOT ANSWERED 9 OUT OF UNIVERSE	1 111 00099 00000	CODE 0

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STREET OR HIGHWAY NOISE YES STREET NOISE NOISTREET NOISE NOISE. BOTHERSOME NOISE. NO BOTHER OBJECTS. WISHES TO MOVE OBJECTS. DOES NOT WISH TO MOVE B NOT ANSWERED		NO. OF CARS IN HOUSEHOLD NO CARS 1 CAR 2 CARS 1 3 CARS 1 4 DR MORE CARS	OWN SECOND HOME POWN SECOND HOME	PARKING SPACE NOT CONNECTED WITH BLDG YES ANOTHER LOCATION NO NOT ANSWERED OUT OF UNIVERSE	RENT INCLUDES PARKING COST YES INCLUDED IN RENT NO PAID SEPARATELY NOT ANSWERED OUT OF UNIVERSE	MONTHLY COST OF PARKING \$199 OR MORE NOT ANSWERED OUT OF UNIVERSE	PARKING SPACE RENTED IN BUILDING YES PARKING RENTED NO PARKING NOT RENTED OUT OF UNIVERSE	DATA SECTION DESCRIPTIVE LABELS	DATA BASE DICTIONARY
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	3 TRAFFIC. BOTHERSOME 4 TRAFFIC. NO BOTHER 5 OBJECTS. WISHES TO MOVE 6 OBJECTS. DOES NOT WISH TO MOVE 8 NOT ANSWERED			
H209	STREETS NEED REPAIR 1 YES STREETS NEED REPAIR 2 NO. DO NOT NEED REPAIR 3 NEED REPAIR. BOTHERSOME 4 NEED REPAIR. NO BOTHER 5 OBJECTS. WISHES TO MOVE 6 OBJECTS. DOES NOT WISH TO MOVE 8 NOT ANSWERED		123 00099 00000	CODE
Н210	STREETS IMPASSABLE DUE TO SNOW, ETC 1 YES, STRIETS IMPASSABLE 2 NO. NOT IMPASSABLE 3 IMPASSABLE, BOTHERSOME 4 IMPASSABLE, NO BOTHER 5 OBJECTS, WISHES TO MOVE 6 OBJECTS, DOES NOT WISH TO MOVE 8 NOT ANSWIRED		124 00099 00000	CODE
H211	INADEQUATE STREET LIGHTING YES. INADEQUATE LIGHTING NO INADEQUATE LIGHTING INADEQUATE LIGHTING. BOTHERSOME INADEQUATE LIGHTING. NO BOTHER BUBLECTS. WISHES TO MOVE BUBLECTS. DOES NOT WISH TO MOVE	-	125 00099 00000	CODE
ភ :	NEIGHBORHOOD CRIME 1 YES. NEIGHBORHOOD CRIME 2 NO NEIGHBORHOOD CRIME 3 CRIME, BOTHERSOME 4 CRIME, NO BOTHER 5 OBJECTS, WISHES TO MOVE 6 OBJECTS, DOES NOT WISH TO MOVE 8 NOT ANSWERED		126 00099 00000	CODE
H217	ODORS OR SMOKE 1 YES. ODORS, SMOKE, OR GAS 2 NO ODORS 3 ODORS, BOTHERSOME 4 ODORS, NO BOTHER 5 OBJECTS, WISHES TO MOVE 6 OBJECTS, DOES NOT WISH TO MOVE 8 NOT ANSWERED		127 00099 00000	CODE

H225	H224	H223	H222	H221	H220	H219	H218	NAME		
ADEQUATE POLICE PROTECTION 1 YES. ADEQUATE 2 NO. INADEQUATE 3 DONT KNOW 8 NOT ANSWERED	INADEQUATE SHOPPING- WISH TO MOVE 1 YES. WISHES TO MOVE 2 NO. DOES NOT WISH TO MOVE 8 NOT ANSWIRED 9 OUT OF UNIVERSE	ADEQUATE SHOPPING 1 YES ADEQUATE 2 NO. INAD! QUATE 3 DONT KNOW 8 NOT ANSWIRED	INADEQUATE SCHOOLS- WISH TO MOVE 1 YES, WISHES TO MOVE 2 NO. DOES NOT WISH TO MOVE 8 NOT ANSWIRED 9 OUT OF UNIVERSE	ADEQUATE SCHOOLS 1 YES ADEQUATE 2 NO. INADEQUATE 3 DONT KNOW 8 NOT ANSWERED	INADEQUATE TRANSPORTATION: WISH TO MOVE 1 YES, WISHES TO MOVE 2 NO. DOES NOT WISH TO MOVE 8 NOT ANSWIRED 9 OUT OF UNIVERSE	ADEQUATE PUBLIC TRANSPORTATION 1 YES. ADEQUATE 2 NO. INADEQUATE 3 DONT KNOW 8 NOT ANSWERED	AIRPLANE NOISE 1 YES AIRPLANE NOISE 2 NO AIRPLANE NOISE 3 NOISE, BOTHERSOME 4 NOISE, BOTHER 5 OBJECTS, WISHES TO MOVE 6 OBJECTS, DOES NOT WISH TO MOVE 8 NOT ANSWERED	DESCRIPTIVE LABELS	DATA SECTION	DATA BASE DICTIONARY
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H244	H243	H242	H241	H240	H231	H230	H229	H228	H227	H2 26	NAME	
LINE NO. OF PERSON 3 99 OUT OF UNIVERSE	WAGE SALARY INCOME- PERSON 2	Y LINE NO. OF PERSON 2	WAGE SALARY INCOME- PERSON 1	LINE NUMBER OF FAMILY MEMBER 1 99 OUT OF UNIVERSE	NEIGHBORHOOD RATING 1 EXCELLENT 2 GOOD 3 FAIR 4 POOR 8 NOT ANSWERED	INADEQUATE HEALTH CLINICS 1 YES, WISHES TO MOVE 2 NO. DOES NOT WISH TO MOVE 8 NOT ANSWERED 9 OUT OF UNIVERSE	ADEQUATE HEALTH CLINICS 1 YES. ADEQUATE 2 NO. INADEQUATE 3 DONT KNOW 8 NOT ANSWERED	INADEQUATE FIRE PROTECTION 1 YES. WISHES TO MOVE 2 NO. DOES NOT WISH TO MOVE 8 NOT ANSWERED 9 OUT OF UNIVERSE	ADEQUATE FIRE PROTECTION 1 YES. ADEQUATE 2 NO. INADFQUATE 3 DONT KNOW 8 NOT ANSWERED	INADEQUATE POLICE PROTECTION 1 YES. WISHES TO MOVE 2 NO. DOES NOT WISH TO MOVE 8 NOT ANSWIRED 9 OUT OF UNIVERSE	DATA SECTION DESCRIPTIVE LABELS	DATA BASE DICTIONARY
2 156 00999 00000	5 151	2 149 00999 00000	5 144	2 142 00999 00000	1 141 00099 00000	1 140 00099 00000	1 139 00099 00000	1 138 00099 00000	1 137 00099 00000	1 136 00099 00000	LENGTH BEGIN MAX. MIN.	c3/20/79
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OF SOCIAL SECURITY INCOME AMOUNT - WELFARE INCOME AMOUNT - WELFARE INCOME AMOUNT - WELFARE INCOME AMOUNT - PRIVATE PENSION AMOUNT - OTHER INCOME SOURCES 14 YEARS OLD AND OVER WITH WAGE INCOME 15 206 16 195 NUMBER 17 WAGE AND SALARY INCOME 18 201 NUMBER 20 201	LINE NO. OF PERSON 5 PAGE SALARY INCOME - PERSON 5 LINE NO. OF FERSON 5 LINE NO. OF F	WAGE SALARY INCOME - PERSON 4 5 165 NUMBER	LINE NO. OF PERSON 4 WAGE SALARY INCOME - PERSON 4 LINE NO. OF FERSON 5 LINE NO. OF FERSON 6 IF MORE THAN 6 PERSON 6 IF MORE THAN 6 PERSON 6 IF FOR THE REMAINING PERSON 6 WAGE SALARY INCOME - P	WAGE SALARY INCOME - PERSON 3	MAME DESCRIPTIVE LABELS WAGE SALARY INCOME. PERSON 3 LINE NO. OF PERSON 4 99 OUT OF UNIVERSE WAGE SALARY INCOME. PERSON 4 LINE NO. OF FERSON 5 99 OUT OF UNIVERSE WAGE SALARY INCOME. PERSON 5 LINE NO. OF FERSON 5 99 OUT OF UNIVERSE WAGE SALARY INCOME. PERSON 5 LINE NO. OF FERSON 6 LINE NO. 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HIGHEST GRADE HEAD COMPLETED ON NEVER ATTENDED SCHOOL O1 KINDERGARTEN O2 FIRST O3 SECOND O4 THIRD O5 FOURTH O6 FIFTH O7 SIXTH O8 SEVENTH	ETHNIC ORIGIN OF HEAD O1 MEXICAN-AMERICAN O2 CHICANO O3 MEXICANO O4 MEXICANO O5 PUERTO RICAN O6 CUBAN O7 CENT OR SOUTH ATER O8 OTHER SPANISH O9 OTHER	HEADS WORK DISTANCE [ONE WAY] 000 LESS THAN 1 MILE 201 201 MILES OR MORE 999 OUT OF UNIVERSE	HEADS TIME FROM HOME TO WORK 199 199 MINUTES OR MORE 999 OUT OF UNIVERSE	101 TRUCK. DRIVES ALONE 102 TRUCK. SHARES DRIVING 103 TRUCK. DRIVES OTHERS 104 TRUCK. RIDES WITH SOMEONE ELSE 201 AUTO. DRIVES ALONE 202 AUTO. SHARES DRIVING 203 AUTO. DRIVES OTHERS 204 AUTO. RIDES WITH SOMEONE ELSE 005 WALKS ONLY 006 WORKS AT HOME 007 RAILROAD 009 BUS OR STREETCAR 010 TAXICAB 011 MOTORCYCLE 012 OTHER MEANS 013 BICYCLE 014 DE UNIVERSE	m	DESCRIPTIVE LABELS	DATA BASE DICTIONARY
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	06	TO ATTEND SCHOOL						
	80	NEEDED LARGER HOUSE OR APT		•				
	1 09	SEPARATED						
	.	DIVORCED MOVED NEARER RELATIVES						
	1 3	NEWLY MARRIED .						
	- - -	FAMILY DECREASED WANTED OWN HOUSEHOLD						
	17	NEIGHEGHOOD OVERGROWNED						
	19	RACIAL/ETHNIC CHANGE IN NEIGHBORHOOD						
	210	WANTED TO OWN RESIDENCE						
	2 22	LOWER RENT OR LESS EXPENSIVE HOUSE						
	24	BY URBAN RE						
	26	DISPLACED BY PRIVATE ACTION						

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RELATIONSHIP TO HEAD- OTHER RELATIVE PARTNER OTHER NON-RELATIVE WIFE UNMARRIED CHILD HEAD [SUBFAMILY 1] CHILD [SUBFAMILY 1] HEAD [SUBFAMILY 2] HEAD [SUBFAMILY 2] WIFE [SUBFAMILY 2] WIFE [SUBFAMILY 3] WIFE [SUBFAMILY 3] CHILD [SUBFAMILY 3] HEAD [SUBFAMILY 3] WIFE [SUBFAMILY 3] HEAD [SUBFAMILY 4] WIFE [SUBFAMILY 4]	LINE NO. OF PERSON 02 OUT OF UNIVERSE	SEX OF HEAD MALE FEMALE	RACE OF HEAD WHITE -BLACK OTHER	MARITAL STATUS OF HEAD MARRIED WIDOWED DIVORCED SEPARATED NEVER MARRIED	AGE OF HEAD 97 YEARS AND OVER	RELATIONSHIP TO HEAD HEAD	WANTED TO RENT MORE CONVENIENCES NATURAL DISASTER CHANGE CLIMATE OTHER NOT ANSWERED OUT OF UNIVERSE	DATA SECTION DESCRIPTIVE LABELS	DATA BASE DICTI
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AGE03		REL03	CONN	SEX02	RACOZ	WST02	AGEO2	NAME	
AGE OF PERSON 03	O5 PARTNER O6 OTHER NON-RELATIVE 12 WIFE 13 UNMARRIED CHILD 21 HEAD [SUBFAMILY 1] 22 WIFE [SUBFAMILY 1] 23 CHILD [SUBFAMILY 2] 34 WIFE [SUBFAMILY 2] 35 CHILD [SUBFAMILY 2] 36 CHILD [SUBFAMILY 2] 37 WIFE [SUBFAMILY 2] 38 WIFE [SUBFAMILY 2] 39 WIFE [SUBFAMILY 3] 40 WIFE [SUBFAMILY 3] 41 HEAD [SUBFAMILY 3] 42 WIFE [SUBFAMILY 3] 43 CHILD [SUBFAMILY 4] 54 WIFE [SUBFAMILY 4] 55 CHILD [SUBFAMILY 4] 56 OUT OF UNIVERSE	RELATIONSHIP TO HEAD- PERSON 03 04 OTHER RELATIVE	LINE NO. OF PERSON 03 99 OUT OF UNIVERSE	SEX OF PERSON 02 1 MALE 2 FEWALE 9 OUT OF UNIVERSE	RACE OF PERSON 02 1 WHITE 2 BLACK 3 OTHER 9 OUT OF UNIVERSE	MARITAL STATUS- PERSON 02 1 MARRIED 2 WIDOWED 3 DIVORCED 4 SEPARATED 5 NEVER MARRIED 9 OUT OF INIVERSE	53 CHILD [SUBFAMILY 4] 99 OUT OF UNIVERSE AGE OF PERSON O2 97 97 YEARS AND OVER 99 OUT OF UNIVERSE	DATA SECTION DESCRIPTIVE LABELS	DATA BASE DICTIONARY
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AGE05	RELOS	LNNOS	SEXO4	RACO4		NAME	
AGE OF PERSON 05 97 97 YEARS AND OVER 99 OUT OF UNIVERSE	RELATIONSHIP TO HEAD. PERSON 05	POUT OF UNIVERSE	SEX DF PERSON 04 1 MALE 2 FEMALE 9 OUT OF UNIVERSE	RACE OF PERSON 04 1 WHITE 2 BLACK 3 OTHER 9 OUT OF UNIVERSE	1 MARRIED 2 WIDOWED 3 DIVORCED 4 SEPARATED 5 NEVER MARRIED 9 OUT OF UNIVERSE	DATA SECTION DESCRIPTIVE LABELS	DATA BASE DICTIONARY
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MST05

MARITAL STATUS OF PERSON 05
MARRIED
WIDOWED
DIVORCED
SEPARATED

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RACO6	MST06	AGE06		RELOG	LNN06	SEX05	RACOS		NAME	
RACE OF PERSON OF	MARITAL STATUS OF PERSON D6 1 MARRIED 2 WIDOWED 3 DIVORCED 4 SEPARATED 5 NEVER MARRIED 9 OUT OF UNIVERSE	AGE OF PERSON OG 97 97 YEARS AND OVER 99 OUT OF UNIVERSE	OF PARTNER OF OTHER NON-RELATIVE 12 WIFE 13 UNMARRICD CHILD 21 HEAD [SUBFAMILY 1] 22 WIFE [SUBFAMILY 1] 23 CHILD [SUBFAMILY 2] 34 HEAD [SUBFAMILY 2] 35 CHILD [SUBFAMILY 2] 46 HEAD [SUBFAMILY 3] 47 WIFE [SUBFAMILY 3] 48 CHILD [SUBFAMILY 3] 59 WIFE [SUBFAMILY 4] 50 CHILD [SUBFAMILY 4] 51 HEAD [SUBFAMILY 4] 52 WIFE [SUBFAMILY 4] 53 CHILD [SUBFAMILY 4] 54 WIFE [SUBFAMILY 4] 55 OUT OF UNIVERSE	RELATIONSHIP TO HEAD- PERSON OF	LINE NO. OF PERSON OF	SEX OF PERSON 05 1 MALE 2 FEMALE 9 OUT OF UNIVERSE	RACE OF PERSON OS 1 WHITE 2 BLACK 3 OTHER 9 OUT OF UNIVERSE	5 NEVER MARRIED 9 OUT OF UNIVERSE	DATA SECTION DESCRIPTIVE LABELS	DATA BASE DICTIONARY
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RACO7	MST07	AGEO7		RELO7	LNN07	SEX06		NAME	
RACE OF PERSON 07 1 WHITE 2 BLACK 3 OTHER 9 OUT OF UNIVERSE	MARITAL STATUS OF PERSON 07 1 MARRIED 2 WIDOWED 3 DIVORCED 4 SEPARATED 5 NEVER MARRIED 9 OUT OF UNIVERSE	AGE OF PERSON 07 97 97 YEARS AND OVER 99 OUT OF UNIVERSE	OG OTHER NON-RELATIVE 112 WIFE 13 UNMARRICO CHILD 21 HEAD [SUBFAMILY 1] 22 WIFE [SUBFAMILY 1] 23 CHILD [SUBFAMILY 2] 34 HEAD [SUBFAMILY 2] 35 CHILD [SUBFAMILY 2] 41 HEAD [SUBFAMILY 2] 42 WIFE [SUBFAMILY 3] 43 CHILD [SUBFAMILY 3] 54 WIFE [SUBFAMILY 3] 55 CHILD [SUBFAMILY 4] 56 WIFE [SUBFAMILY 4] 57 WIFE [SUBFAMILY 4] 58 CHILD [SUBFAMILY 4] 59 OUT OF UNIVERSE	23	LINE NO. OF PERSON 07 99 OUT OF UNIVERSE	SEX OF PERSON OF 1 MALE 2 FEMALE 9 OUT OF UNIVERSE	1 WHITE 2 BLACK 3 OTHER 9 OUT OF UNIVERSE	DATA SECTION DESCRIPTIVE LABELS	DATA BASE DICTIONARY
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SEX08	RACO8	MST08	AGEO8		RELOS	LNN08	SEX07	NAME	
SEX OF PERSON OB 1 MALE 2 FEMALE 9 OUT OF UNIVERSE	RACE OF PERSON OB 1 WHITE 2 BLACK 3 OTHER 9 OUT OF UNIVERSE	MARITAL STATUS OF PERSON 08 1 MARRIED 2 WIDOWED 3 DIVORCED 4 SEPARATED 5 NEVER MARRIED 9 OUT OF UNIVERSE	AGE OF PERSON OB 97 97 YEARS AND OVER 99 OUT OF UNIVERSE	OS PARTNER OS OTHER NON-RELATIVE 12 WIFE 13 UNMARRIED CHILD 21 HEAD [SUBFAMILY 1] 22 WIFE [SUBFAMILY 1] 23 CHILD [SUBFAMILY 2] 31 HEAD [SUBFAMILY 2] 32 WIFE [SUBFAMILY 2] 33 CHILD [SUBFAMILY 2] 41 HEAD [SUBFAMILY 3] 42 WIFE [SUBFAMILY 3] 43 CHILD [SUBFAMILY 4] 54 WIFE [SUBFAMILY 4] 55 WIFE [SUBFAMILY 4] 56 CHILD [SUBFAMILY 4] 57 WIFE [SUBFAMILY 4] 58 CHILD [SUBFAMILY 4] 59 OUT OF UNIVERSE	æ	LINE NO. OF PERSON OB 99 OUT OF UNIVERSE	MALE FENALE OUT OF UNIVERS	DATA SECTION DESCRIPTIVE LABELS	DATA BASE DICTIONARY
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LNN09 DESCRIPTIVE LABELS DATA SECTION LENGTH BEGIN MAX. MIN. 339 00999 00000 VALUE DATA IMP.DEC CATEGORY PLACES CODE PAGE 0 29

REL09 99 RELATIONSHIP TO HEAD- PERSON 09
OTHER RELATIVE
PARTNER COUT OF UNIVERSE N 341 00999 00000 CODE 0

OTHER NON-RELATIVE UNNARRIED CHILD

HEAD [SUBFAMILY 1]
WIFE [SUBFAMILY 1]
CHILD [SUBFAMILY 2]
HEAD [SUBFAMILY 2]
WIFE [SUBFAMILY 2]
CHILD [SUBFAMILY 3]
HEAD [SUBFAMILY 3]
CHILD [SUBFAMILY 3]
HEAD [SUBFAMILY 4]
WIFE [SUBFAMILY 4]
OUT OF UNIVERSE

NUMBER 0

N

343

AGE09

9**7**

AGE OF PERSON OS
97 YEARS AND OVER
OUT OF UNIVERSE

MARTIAL STATUS OF PERSON 09

WIDOWED

DIVORCED
SEPARATED
NEVER MARRIED
OUT OF UNIVERSE

WST09

0

0

0

SEX09

SEX OF PERSON 09
MALE

FEMALE OUT OF UNIVERSE

RAC09

RACE OF PERSON 09
WHITE
BLACK

OTHER OUT OF UNIVERSE

REL10

LNN10

99

LINE NO. OF PERSON 10 OUT OF UNIVERSE

RELATIONSHIP TO HEAD- PERSON 10

0

0

N

	REL11	LNN11	SEX10	RAC10	MST 10	AGE 10		NAME	
06 OTHER NON-RELATIVE 12 Wife	RELATIONSHIP TO HEAD- PERSON 11 O4 OTHER RELATIVE O5 PARTNER	LINE NO. OF PERSON 11	SEX OF PERSON 10 1 MALE 2 FEMALE 9 OUT OF UNIVERSE	RACE OF PERSON 10 1 WHITE 2 BLACK 3 OTHER 9 OUT OF UNIVERSE	MARITAL STATUS OF PERSON 10 1 MARRIED 2 WIDOWED 3 DIVORCED 4 SEPARATED 5 NEVER MARRIED 9 OUT OF UNIVERSE	AGE OF PERSON 10 97 97 YEARS AND OVER 99 OUT OF UNIVERSE	O4 OTHER RELATIVE O5 PARTNER O6 OTHER NON-RELATIVE 12 WIFE 13 UNMARRIED CHILD 21 HEAD [SUBFAMILY 1] 22 WIFE [SUBFAMILY 1] 23 CHILD [SUBFAMILY 2] 34 HEAD [SUBFAMILY 2] 35 CHILD [SUBFAMILY 2] 41 HEAD [SUBFAMILY 3] 42 WIFE [SUBFAMILY 3] 43 CHILD [SUBFAMILY 4] 54 WIFE [SUBFAMILY 4] 55 WIFE [SUBFAMILY 4] 56 CHILD [SUBFAMILY 4] 57 WIFE [SUBFAMILY 4] 58 OCHILD [SUBFAMILY 4] 59 OUT OF UNIVERSE	DATA SECTION DESCRIPTIVE LABELS	DATA BASE DICTIONARY
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REC12	LNN12	SEX11	RAC11	#S711	AGE11 9	9 5 5 5 4 4 4 3 3 3 5 5 5 5 5 5 5 5 5 5 5	NAME	
RELATIONSHIP TO HEAD- PERSON 12 O4 OTHER RELATIVE O5 PARTNER O6 OTHER NON-RELATIVE 12 WIFE 13 UNMARRIED CHILD	LINE NO. OF PERSON 12 99 OUT OF UNIVERSE	SEX OF PERSON 11 1 MALE 2 FEMALE 9 OUT OF UNIVERSE	RACE OF PERSON 11 1 WHITE 2 BLACK 3 OTHER 9 OUT OF UNIVERSE	MARTIAL STATUS OF PERSON 11 1 MARRIED 2 WIDOWED 3 DIVORCED 4 SEPARATED 5 NEVER MARRIED 9 OUT OF UNIVERSE	AGE OF PERSON 11 97 97 YEARS AND OVER 99 OUT OF UNIVERSE	UNMARRIE HEAD [SU WIFE [SU CHILD [SU WIFE [SU CHILD [SU WIFE [SU WIFE [SU WIFE [SU WIFE [SU WIFE [SU WIFE [SU	DATA SECTION DESCRIPTIVE LABELS	DATA BASE DICTIONARY
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REL 13	LNN13	SEX12	RAC12	#S712	AGE 12		NAME	
RELATIONSHIP TO HEAD- PERSON 13 O4 OTHER RELATIVE O5 PARTNER O6 OTHER NON-RELATIVE 12 WIFE 13 UNMARRIED CHILD 21 HEAD [SUBFAMILY 1] 22 WIFE [SUBFAMILY 1] 23 CHILD [SUBFAMILY 2] 34 HEAD [SUBFAMILY 2] 35 WIFE [SUBFAMILY 2] 36 CHILD [SUBFAMILY 2] 37 WIFE [SUBFAMILY 3]	LINE NO. OF PERSON 13 99 OUT OF UNIVERSE	SEX OF PERSON 12 1 MALE 2 FEMALE 9 OUT OF UNIVERSE	RACE OF PERSON 12 1 WHITE 2 BLACK 3 OTHER 9 OUT OF UNIVERSE	MARITAL STATUS OF PERSON 12 1 MARRIED 2 WIDOWED 3 DIVORCED 4 SEPARATED 5 NEVER MARRIED 9 OUT OF UNIVERSE	AGE OF PERSON 12 97 97 YEARS AND OVER 99 OUT OF UNIVERSE	31 HEAD [SUBFAMILY 2] 32 WIFE [SUBFAMILY 2] 33 CHILD [SUBFAMILY 2] 41 HEAD [SUBFAMILY 3] 42 WIFE [SUBFAMILY 3] 53 CHILD [SUBFAMILY 4] 54 WIFE [SUBFAMILY 4] 55 CHILD [SUBFAMILY 4] 56 CHILD [SUBFAMILY 4] 57 CHILD [SUBFAMILY 4] 58 CHILD [SUBFAMILY 4]	DATA SECTION DESCRIPTIVE LABELS	DATA BASE DICTIONARY
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	REL14	NN14	SEX13	RAC13	MST13	AGE 13		NAME	
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OTHER RELATIVE PARTNER OTHER NON-RELATIVE WIFE UNMARRIED CHILD HEAD [SUBFAMILY 1] CHILD [SUBFAMILY 1] HEAD [SUBFAMILY 2] WIFE [SUBFAMILY 2] WIFE [SUBFAMILY 2] HEAD [SUBFAMILY 3] WIFE [SUBFAMILY 3] WIFE [SUBFAMILY 3] CHILD [SUBFAMILY 4] WIFE [SUBFAMILY 4]	RELATIONSHIP TO HEAD- PERSON 14	LINE NO. OF PERSON 14 OUT OF UNIVERSE	SEX OF PERSON 13 MALE FEMALE OUT OF UNIVERSE	RACE OF PERSON 13 WHITE BLACK OTHER OUT OF UNIVERSE	MARITAL STATUS OF PERSON 13 MARRIED WIDOWED DIVORCED SEPARATED NEVER MARRIED OUT OF UNIVERSE	AGE OF PERSON 13 97 YEARS AND OVER OUT OF UNIVERSE	WIFE [SUBFAMILY 3] CHILD [SUBFAMILY 3] HEAD [SUBFAMILY 4] WIFE [SUBFAMILY 4] CHILD [SUBFAMILY 4] OUT OF UNIVERSE	DATA SECTION DESCRIPTIVE LABELS	DATA BASE DICTIONARY
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AGE 15		REL15	LNN15	SEX14	RAC14	MST 14	AGE 14	NAME	
AGE OF PERSON 15	OF PARTNER OF OTHER NON-RELATIVE 12 WIFE 13 UNMARRIED CHILD 21 HEAD [SUBFAMILY 1] 22 WIFE [SUBFAMILY 1] 23 CHILD [SUBFAMILY 2] 31 HEAD [SUBFAMILY 2] 32 WIFE [SUBFAMILY 2] 33 CHILD [SUBFAMILY 3] 44 HEAD [SUBFAMILY 3] 45 WIFE [SUBFAMILY 3] 56 WIFE [SUBFAMILY 3] 57 WIFE [SUBFAMILY 4] 58 CHILD [SUBFAMILY 4] 59 OUT OF UNIVERSE	RELATIONSHIP TO HEAD- PERSON 15	LINE NO. OF PERSON 15	SEX OF PERSON 14 1 MALE 2 FEMALE 9 OUT OF UNIVERSE	RACE OF PERSON 14 1 WHITE 2 BLACK 3 OTHER 9 OUT OF UNIVERSE	MARITAL STATUS OF PERSON 14 1 MARRIED 2 WIDOWED 3 DIVORCED 4 SEPARATED 5 NEVER MARRIED 9 OUT OF UNIVERSE	53 CHILD [SUBFAMILY 4] 99 OUT OF UNIVERSE AGE OF PERSON 14 97 97 YEARS AND OVER 99 OUT OF UNIVERSE	DESCRIPTIVE LABELS	DATA BASE DICTIONARY
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397		395 00999 00000	393 00999 00000	392 00099 00000	391 00099 00000	390 00099 00000	388	VALUE LENGTH BEGIN MAX. MIN.	03/20/79
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HIMEM	SEX15	RAC15	SST15		NAME	
NUMBER OF PERSONS IN HOUSEHOLD 19 19 OR MORE PERSONS	SEX OF PERSON 15 1 MALE 2 FEMALE 9 OUT OF UNIVERSE	RACE OF PERSON 15 1 WHITE 2 BLACK 3 OTHER 9 OUT OF UNIVERSE	MARITAL STATUS OF PERSON 15 1 MARRIED 2 WIDOWED 3 DIVORCED 4 SEPARATED 5 NEVER MARRIED 9 OUT OF UNIVERSE	97 97 YEARS AND OVER 99 OUT OF UNIVERSE	DATA SECTION DESCRIPTIVE LABELS	DATA BASE DICTIONARY
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SOURCE AND RELIABILITY STATEMENT

SAMPLE DESIGN

The DOT Travel-to-Work Supplement and the Annual Housing Survey

The DOT Travel-to-Work Supplement data are based on interviews completed during the period April 1975 through March 1976 in 21 SMSA's as part of the enumeration for the Year II Annual Housing Survey (AHS) sponsored by the Department of Housing and Urban Development. In the largest SMSA from each of the four geographic regions, the survey sample consisted of about 15,000 housing units evenly divided between the central city or cities and the balance of the SMSA. For the remaining 17 SMSA's, the survey was based on a sample of about 5,000 housing units distributed between the central city or cities and the balance of the SMSA in proportion to the distribution of total housing units in each of these sectors.

The following SMSA's, defined as of the 1970 census, were included in the 1975-1976 AHS enumeration:

Atlanta, Ga. *
Chicago, Ill. *
Cincinnati, Ohio-Ky.-Ind.
Colorado Springs, Colo.
Columbus, Ohio
Hartford, Conn.
Kansas City, Mo.-Kan.
Miami, Fla.
Milwaukee, Wis.
New Orleans, La.
Newport News-Hampton, Va.

Paterson-Clifton-Passaic, N.J.
Philadelphia, Pa.-N.J.*
Portland, Oreg.-Wash.
Rochester, N.Y.
San Antonio, Tex.
San Bernardino-Riverside-Ontario, Calif.
San Diego, Calif.
San Francisco-Oakland, Calif. *
Springfield-Chicopee-Holyoke, Mass.-Conn.
Madison, Wis.

*Sample size of 15,000 housing units; all others are 5,000 housing units.

The number of housing units in each SMSA that were eligible for interview in the AHS is given in the first column of table A. Of these sample units, a number of interviews were not obtained (table A, column 2) because, four occupied sample units, the occupants were not at home after repeated visits or were unavailable for some other reason; or, for vacant units, no informed respondent could be found after repeated visits. In addition to units eligible for interview, there were other units that were visited but were found not to be eligible for interview because they were condemned, unfit, demolished, converted to group quarters use, etc., (table A, column 3). The number of persons 14 years and older in the interviewed households of each SMSA are given in the fourth column of table A. The number of persons 14 years and older who did not respond to the DOT Travel-to-Work Supplement is given in the last column of table A.

Table A. Number of Housing Units and Persons 14 Years and Older, Interviewed and Not Interviewed, for 21 SMSA's, 1975-76

	Housing Units Eligible for Interview	Noninterviewed Housing Units	Housing Units Found to be Ineligible for Interview	Persons 14+ in Inter- viewed Housing Units	Persons 14 Not Responding to DO Supplement
				٠	•
SMSA					
Atlanta	13,518	570	1,484	24,612	151
Chicago	14,442	741	1,051	29,117	315
Cincinnati	4,721	173	408	9,491	57
Colorado Springs	4,776	135	428	8,866	140
Columbus	4,749	203	464	9,028	65
Hartford	4,926	213	322	10,337	41
Kansas City	4,684	210	472	8,848	44
Madison	4,899	129	495	9,756	17
Miami	4,537	216	473	8,248	21
Milwaukee	4,931	204	298	10,317	38
New Orleans	4,746	231	422	9,104	45
Newport News-	4,105	118	492	8,235	19
Paterson-	4,981	253	193	10,976	39 -
Philadelphia	14,447	863	1,203	29,397	207
Portland	4,875	157	363	9,264	46
Rochester	4,845	205	508	9,843	26
San Antonio	4,883	170	421	9,802	26
San Bernardino-	4,963	241	347	8,515	36
San Diego	4,556	212	352	8,542	68
San Francisco-	14,523	861	934	26,135	359
Springfield-	5,239	167	504	10,629	40

Selection of the AHS-SMSA sample The sample for the SMSA's which are 100 percent permit issuing was selected from two sample frames: units enumerated in the 1970 Census of Population and Housing in areas under the jurisdiction of permit-issuing offices (the permit-issuing universe) and units constructed in permit-issuing areas since the 1970 census (the new construction universe). In addition, the sample for those SMSA's which are not 100 percent permit-issuing included a sample selected from a third frame--those units located in areas not under the jurisdiction of permit-issuing offices (the non-permit universe). The following table indicates which SMSA's are 100 percent issuing and which contain a sample from the non-permit universe.

100% Permit-Issuing SMSA's

Hartford, Conn.
Miami, Fla.
Newport News-Hampton, Va.
Paterson-Clifton-Passaic,
N.J.
San Bernardino-RiversideOntario, Calif.
San Diego, Calif.
San Francisco-Oakland,
Calif.

SMSA's with a Sample from the Non-Permit Universe

Atlanta, Ga.
Chicago, Ill.
Cincinnati, Ohio-Ky.-Ind.
Colorado Springs, Colo.
Columbus, Ohio
Kansas City, Mo.-Kan.
Madison, Wis.
Milwaukee, Wis.
New Orleans, La.
Philadelphia, Pa.-N.J.
Portland, Oreg.-Wash.
Rochester, N.Y.
San Antonio, Tex.
Springfield-Chicopee-Holyoke,
Mass.-Conn.

The major portion of the sample in each SMSA was selected from a file which represented the 20-percent sample of units enumerated in permit-issuing areas of the SMSA during the 1970 Census of Housing and Population. This file contained records for occupied housing units, vacant housing units, and units in certain special places or group quarters. Sampling operations were done separately for the special place and group quarters records and for the occupied and vacant housing unit records. Before the sample was selected from the occupied and vacant records, the occupied housing unit records were stratified by race of head (non-Black/Black) and the vacant records were stratified into four categories pertaining to the value or rent associated with the vacant housing units. The occupied housing unit records were further stratified such that each unit was assigned to one of 50 strata according to its tenure/family size/household income category as illustrated by the following table:

** * * * *	L				Ten	ure			
Household			Own				R	ent	er
Income		Fan	ily	Si	ze				Size
	1	2	3	4	5+	1	2	3	4 5+
Under \$3,000 \$3,000-\$5,999 \$6,000-\$9,999 \$10,000-\$14,999 \$15,000 and Over					•				

Thus, for each SMSA, the occupied housing unit records from this universe were assigned to one of 100 strata for either the central city or for the balance, and the vacant housing unit records were assigned to one of the four vacant strata for either the central city or for the balance of the SMSA. A sample of housing unit records was then selected in such a way as to produce one-half of the desired sample size. However, the housing unit record adjacent to each of the above sample housing unit records was also selected to be in sample thereby insuring the necessary designated sample size.

Before the sample was selected from the group quarters and special place records, the records were stratified by census tract and census enumeration district (ED) within the central city and within the balance of the SMSA. A sample of special place records was then selected in such a way as to produce one-quarter of the desired sample size. However, at the time of enumeration the units at each of the sample special places were listed and subsampled at a rate so as to produce an expected four sample units, thereby insuring the necessary designated sample size.

The second frame from which each SMSA sample was selected was a list of new construction building permits issued since 1970 (i.e., the new construction universe). The sample selection from the list of new construction building permits was an independent operation within this SMSA. Prior to sample selection, the list of permits was chronologically stratified by the date the permits were issued, and clusters of an expected four (usually adjacent) housing units were formed. These clusters were then sampled for inclusion at the overall sampling rate.

For those SMSA's which are not 100 percent permit-issuing, the remainder of the AHS sample was selected from a frame consisting of areas not under the jurisdiction of permit-issuing offices (i.e., the non-permit universe). The first step in the sampling operation for the universe was the selection, using the overall sampling rate, of a sample of census enumeration districts

within these areas. Prior to this sample selection, the ED's were stratified by census tract within the central city and within the balance of the SMSA. The probability of selection of an ED was proprotionate to the following measure of size:

Number of housing units in 1970 census ED	n +	Group quarters population 1970 census ED	in
		3	
	A		

The sample ED's were then divided into segments, i.e., small land areas with well-defined boundaries having an expected size of four, or a multiple of four, housing units. At the time of enumeration, those segments that did not have an expected size of four were further subdivided to produce an expected four sample housing units.

The next step was the selection of one of these segments within each sample ED. In the sample segments, all units in existence at the time of interview are in sample. Thus, units enumerated in the 1970 census as well as units built since the 1970 census are included.

ESTIMATION

AHS-SMSA Housing Inventory Estimation Procedure

The AHS estimates of characteristics of the 1975 housing inventory employed a three-stage ratio estimation procedure. However, the third stage was employed in only three SMSA's. Initially, the basic weight (i.e., the inverse of the probability of selection) for each interviewed sample housing unit was adjusted to account for the noninterviews previously mentioned. The noninterview adjustment factor was equal to the following ratio:

Weighted count of interviewed housing units	+	hted count of noninterviewed housing units
Weighted count of	interviewed h	housing units

Within each sector (central city and balance) of each SMSA, a noninterview factor was computed separately for 54 noninterview cells for sample housing units from the permit-issuing universe (where the cells consisted of one or more of the different strata used in the stratification of the universe as previously illustrated), one noninterview cell for new construction sample housing units and one noninterview cell for the non-permit universe (if applicable). Sample housing units from the non-permit universe identified as being built April 1, 1970 or later were considered as new construction units for the purpose of the noninterview adjustment.

The first-stage ratio estimation procedure was then employed for all sample housing units from the permit-issuing universe. This factor was computed separately for all sample housing units within each permit-issuing universe noninterview cell mentioned above. The ratio sermit-issuing universe noninterview cell mentioned above. The ratio sermit-issuing universe noninterview cell mas equal to the following:

1970 census count of housing units from permit-issuing universe in a cell AHS sample estimate of 1970 housing units from the cell

For each SMSA the numerators of the ratios were obtained from the 1970 Cansus of Population and Housing 20-percent file of units enumerated in areas under the jurisdiction of permit-issuing offices. The denominators of the ratios were obtained from weighted estimates of all AHS sample units within the corresponding ratio estimate categories using the existing weight (i.e., the basic weight times the noninterview factor). The computed first-stage ratio estimation factor was then applied to the existing weight for each sample unit within the corresponding first-stage ratio estimation category.

The effect of this ratio estimation procedure was to somewhat reduce the variation in sample size for strata used in the sample selection of the permit-issuing universe. Ordinarily, this would have been controlled by sampling within the strata during the sample selection process. However, prior to the AHS sample selection within each SMSA, units already selected for other Census Bureau surveys were deleted from the permit-issuing universe. Thus, some variation in sample size was introduced during the AHS sample selection process.

The second-stage ratio estimation procedure was employed for all sample units that were conventional new construction units (i.e., mobile homes units that were conventional new mobile homes (i.e., mobile homes built after April 1970), new mobile homes (i.e., units added by conversion of 1970 units or by other sources). This procedure was designed to adjust the AHS sample estimates of these types of units to independently derived current estimates. These independent estimates units. This adjustment was necessary to correct for those types of units. This adjustment was necessary to correct for known deficiencies in the AHS sample with regard to representation of these units (see the section on nonsampling error). The second-stage ratio estimate factors were computed separately for each of the cells in the estimate factors were computed separately for each of the cells in the estimate factors were computed separately for each of the cells in the estimate factors were computed separately for each of the cells in the

October 1975 independent estimate of housing units in that cell for the SMSA

ANS sample estimates of housing units in that cell

ASMS and Tor

Second-Stage Ratio Estimate Cella
Conventional New Construction Units

Wew Mobile Hones

"Other Additions"

The numerators of the ratios were derived by applying the following factors: (1) For the "conventional new construction units" cell, either a national trend for missed conventional new construction was applied to the 1968 and 1969 building permits issued in the SMSA, or the missed conventional new construction was estimated from the Survey of Construction (SOC). (The table that follows this paragraph indicates the procedure used to estimate the missing conventional new construction for each SMSA.); (2) for the "new mobile homes" cell, a 1970 census relationship between total new construction and new mobile homes that existed in this SMSA for the 1965-1970 period was applied; and

(3) for the "other additions" cell, the "other additions" rate from SCARF was applied.

SMSA's Employing the SOC Estimate
of Missed Conventional
New Construction

Atlanta, Ga.
Chicago, Ill.
Cincinnati, Chio-Ky.-Ind.
Columbus, Chio
Kansas City, Mo.-Kan.
Miami, Fla.
Milwaukee, Wis.
Philadelphia, Pa.-N.J.
Portland, Oreg.-Wash.
San Diego, Calif.
San Francisco-Oakland, Calif.

SMSA's Employing the National Trend for Missed Conventional New Construction

Cclorado Springs, Cclo.
Hartford, Conn.
Madison, Wis.
New Orleans, La.
Newport News-Hampton, Va.
Paterson-Clifton-Passaic, N.J.
Rochester, N.Y.
San Antonio, Tex.
San Eernardino-Riverside-Cntric,
Calif.
Springfield-Chicopee-Holyoke,
Mass.-Conn.

The denominators of the ratios were obtained from the weighted estimates for the AHS sample units within each cell, using the existing weight after the first-stage ratio estimation procedure. The computed second-stage ratio estimate factor was then applied to the existing weight for each sample unit in the corresponding second-stage ratio estimation category.

For some Year II SMSA's (i.e., Colorado Springs, Colo.; Miami, Fla. and San Diego, Calif.) a third-stage ratio estimation procedure was also employed. This procedure involved the ratio estimation of the AHS weighted sample estimate of the October 1975 housing inventory to an independent estimate of the SMSA's October 1975 housing inventory. This ratio estimate factor equaled the following:

Independent estimate of the October 1975 housing inventory for the SMSA

AHS sample estimate of the housing inventory

^{1/}SCARF denotes the Survey of Components of Change and Residential Finance, a survey conducted in 1957-1959 by the Census Bureau.

The numerator of the ratio was derived by using the 1970 census estimate of total housing units for the SMSA in conjunction with the estimate of the change in the housing inventory since the census. The estimate of change was based on either administrative records from utility companies, or, when utility data were not available, on estimates of new construction permits and post—census demolition data. The denominator of that ratio was obtained from the weighted estimate of AHS sample units, using the existing weight after the second—stage ratio estimation procedure. The computed third—stage ratio estimation factor was then applied to the existing weight for each sample unit and the resulting product was used as the final weight for tabulation purposes.

The effect of the third-stage ratio estimation procedure, as well as the overall estimation procedure, was to reduce the sampling error for most statistics below what would have been obtained by simply weighting the results of the sample by the inverse of the probability of selection. Since the housing population of the sample differed somewhat by chance from the SMSA as a whole, it can be expected that the sample estimates will be improved when the sample housing population or different portions of it, are brought into agreement with known estimates of the SMSA housing population.

The third-stage ratio estimation procedure was not employed in all Year II SMSA's since the reliability of the independent estimate varied by SMSA, depending on the completeness of the utility data or the permit-demolition data. The independent estimates were only employed in SMSA's where the estimated relative bias was thought to be low enough (i.e., three percent or less over a ten year period) to preclarie the possibility of doing more harm than good to the survey results. A measure of the relative bias of the independent estimate was available based on a comparison between the 1970 census estimate of the April 1970 housing inventory and the April 1970 independent estimate of the housing inventory generated from the procedure used to produce the October 1975 independent estimates. In some SMSA's, the measure of the relative bias of the independent estimate was less than three percent but the independent estimate was still not employed, since ther was evidence that suggested that the procedure for generating the independent estimates was producing substantial overestimates for these STEALS. DOT Supplement Adjustments For the DOT Supplement, the weight resulting from the AHS-SMSA estimation procedure up through the first-stage adjustment described above was adjusted to account for persons in households that were interviewed for AHS-SMSA who did not respond to the travel-to-work section of the questionnaire. This noninterview adjustment factor was calculated separately for each sector of each SMSA. Within each sector of each SMSA, a noninterview factor was computed separately for sex, age, and marital status categories. The following formula was used to calculate the non-interview adjustment factor for each noninterview category:

Weighted estimate of persons in the noninterview category who were interviewed for the DOT Supplement + in households that were interviewed for the AHS-SMSA Weighted estimate of persons in the noninterview category who were not interviewed for the DOT Supplement in households that were interviewed for the AHS-SMSA

Weighted estimate of persons in the noninterview category who were interviewed for the DOT Supplement in households that were interviewed for the AHS-SMSA

The final adjustment for persons interviewed for the DOT Supplement was an additional ratio estimation procedure. This procedure was designed to adjust the AHS-SMSA sample estimate of persons 14 years and older in each SMSA to an independently derived current estimate of that same population group. In SMSA's where there was no evidence of differential undercoverage of persons within the sectors, the sample estimate of persons 14+ in the SMSA was adjusted to an independently derived estimate of persons 14+ in the SMSA. For SMSA's where there was evidence of differential undercoverage within the sector, this ratio estimation was performed separately by central city and balance of the SMSA. The factor used for the ratio estimation procedure was calculated as follows:

Independent estimate of persons 14+ in the SMSA (or sector)

Sample estimate of persons 14+ in the SMSA (or sector)

The numerator of this ratio was based on the Census Bureau's estimates of population 14+ as of October 1, 1975. The denominator of this ratio was obtained from the weighted estimate of persons interviewed for the DOT Supplement, using the existing weight after the DOT Supplement nonincerview adjustment had been applied.

The following table shows which person ratio estimate formula was used for each SMSA:

SMSA's for which the person ratio estimate factors were calculated separately by central city and balance

SMSA's for which the person ratio estimate factor was calculated for the entire SMSA

Atlanta, Ga.
Cincinnati, Ohio-Ky.-Ind.
Chicago, III.
Madison, Wis.
Miami, Fla.
Milwaukee, Wis.
New Orleans, La.
Paterson-Clifton-Passaic,
N.J.
Philadelphia, Pa.-N.J.
Rochester, N.Y.
San Diego, Calif.
Springfield-Chicopee-Holyoke,
Mass.-Conn.

Columbus, Ohio
Colorado Springs, Colo.
Hartford, Conn.
Kansas City, Mo.-Kans.
Newport News-Hampton, Va.
Portland, Oreg.-Wash.
San Antonio, Tex.
San Bernardino-RiversideOntario, Calif.
San Francisco-Oakland, Calif.

The computed ratio estimate was then applied to the existing weight for each person interviewed for the DOT Supplement. The weight that resulted from the application of this final adjustment was the tabulation weight utilized to produce final tabulations.

The effect of this person ratio estimation, as well as the overall estimation procedure, was to reduce the sampling error for most statistics below what would have been obtained by simply weighting the results of the sample by the inverse of the probability of selection. Since the population 14 years and older of the sample differed somewhat by chance from the actual population in each city, SMSA balance, or SMSA as a whole, it can be expected that the sample estimates will be improved when the sample population is brought into agreement with known independent estimates of the actual population.

Also, in sample survey enumeration, as in a complete census, certain segments of the population tend to be underrepresented. This underrepresentation results from the fact that the sample frame was incomplete (see section on nonsampling error) and that it was not possible to identify and/or interview all people who should have been interviewed within sample housing units. Methods employed in this estimation procedure, the ratio estimations and noninterview adjustments, are attempts to adjust for this underrepresentation by assuming that the travel-to-work characteristics of persons missed during enumeration or persons who were noninterviews for this survey are similar to the characteristics of persons with the same demographic characteristics who were interviewed.

RELIABILITY OF THE DATA

There are two types of possible errors associated with data from sample surveys: sampling and nonsampling errors. The following is a description of the sampling and nonsampling errors associated with the DOT Travel-to-Work Supplement and AHS—SMSA estimates.

Nonsampling Errors

In general, nonsampling errors can be attributed to many sources: inability to obtain information about all cases, definitional difficulties, differences in the interpretation of questions, inability or unwillingness to provide correct information on the part of respondents, mistakes in recording or coding the data, and other errors of collection, response, processing, coverage, and estimation for missing data.

The DOT Travel-to-Work Supplement One possible source of bias in the DOT Travel-to-Work Supplement data is proxy interviewing. That is, responses for a particular worker may have been given by someone else who is not as knowledgeable as the worker himself. For example, the person available for the interview may not know how long it takes the reference person (worker) to travel to work or whether or not the principal means of transportation to work is satisfactory to the worker. Although it is known that biases due to proxy interviewing, as well as other nonsampling errors, could exist in the DOT Travel-to-Work Supplement, the magnitude of these biases is unknown.

Reinterview Program No reinterview program was undertaken for the DOT Travel-to-Work Supplement. However, for the 1975 AHS-SMSA sample, a study was conducted to obtain a measurement of some of the components of the nonsampling error associated with the AHS estimates. Results of this study may be a useful indicator of the accuracy to be expected in the travel-to-work data which was collected as a supplement to the AHS-SMSA data. A reinterview program was conducted for a sample of the AHS households. These households were revisited and answers to some of the questions on the AHS questionnaire were obtained again. The original interview and the reinterview were assumed to be two independent readings and, thus, were the basis for the measurement of the accuracy of the AHS data collected from enumerated households.

As part of the reinterview, a check was made at each of these households to determine if the following was done during the original interview:

- The correct unit was visited.
- The correct number of housing units were interviewed at that address.
- 3. The correct information on 'Year Built" was obtained.
 4. The correct information on 'Tenure' was obtained.

- 5. The correct information on 'Household Composition' was obtained.
 6. The correct information on 'Type of Housing Unit' was obtained.
- The correct information on 'Occupancy Status' was obtained.

This check was made for interviewer evaluation and control. That is, tolerance limits were derived to determine which interviewers passed or failed this reinterview with regard to the above items. The results of this study are presented in the following Census Bureau memorandum, "Response Error in the Annual Housing Survey Data - Year II SMSA Sample".

Some of the results of this study are presented below (note that these results are based on the reinterviews across all the SMSA's that were part of the 1975-76 sample and not specifically for any SMSA):

- "The results indicate that the interviewers are doing a good job. From a total of 207 interviewers checked, ten failed in classification checks, seven in household composition, and two in content reinterview".
- 2. 'Of the non-attitudinal items which had enough data to compute reliable measures of response error, approximately 55 percent of these items had low levels of inconsistency in responses and about 45 percent had moderate levels of inconsistency in responses."
- 3. 'For the attitudinal items, approximately 60 percent of the indices showed low levels of inconsistency when reconciled, i.e., after the question is answered in the reinterview, the enumerator presents the previous response and then asks the respondent to decide upon the best answer. The remaining items showed moderate levels of inconsistent response."
- 'When no reconciliation was done for these six attitudinal items, about half of the items showed moderate levels of inconsistency, 30 percent were in the low level and 20 percent were in the high inconsistency level."

The range for evaluating inconsistency is from 0-100. "The rule of thumb is that indices below 20 are low, indices from 20-50 are moderate. indicating that there is some problem with inconsistent reporting, and those over 50 are high, indicating that improvements are needed in the method used to collect these data or that the category concepts themselves are ambiguous.

The results of this study were based on sample data so there is sampling error associated with these estimates of nonsampling error. Therefore, the possibility of such errors should be taken into account when considering the results of this study.

With respect to errors of coverage and estimation for missing data, it is believed that the AHS new construction sample had deficiencies with regard to the representation of both conventional new construction and new mobile homes (and trailers) in permit-issuing areas. During the sampling of building permits, only those issued January 1, 1970 or later were eligible to be sampled to represent conventional new construction in permit-issuing areas in the SMSA's. It had been assumed that units with permits issued prior to 1970 would have been completed by the time of the 1970 census (i.e., April 1970), and therefore, would have been represented in the sample selected from 1970 census units. However, it has been estimated that, nationally, there were about 600,000 conventional new construction units built after April 1970 whose permits were issued prior to January 1970 (most of which were issued in 1968 and 1969). Although it is not known exactly, an estimate of the number of conventional new construction units in each SMSA that had permits issued prior to January 1970, and therefore, were missed by the 1975 AHS-SMSA survey, is given in the first column of table B.

In addition, unlike the procedure for conventional new construction, there is no sampling procedure specifically for new mobile homes (and trailers) in permit-issuing areas. However, new mobile homes in permit-issuing areas do come into sample if the mobile homes are located in mobile home parks, identified as such in the 1970 census. Nonetheless, new mobile homes in permit-issuing areas that are located in mobile home parks, not in existence at the time of the 1970 census or not identified as such in the 1970 census, have no chance of coming into the AHS sample. Although it is not known exactly, the estimated number of new mobile homes in permit-issuing areas that were missed by the 1975 AHS-SMSA survey in each SMSA is given in the last column of table B. Finally, it is felt that deficiencies also exist in ED's where area sampling methods are used. As before, it had been assumed that all units located inside these ED's would be represented in the sample. But the 1975 AHS sample has been estimated to miss as much as two percent of all housing units in ED's where area sampling methods are used because these units are not listed during the canvassing.

Therefore, all persons 14 years and older who live in the above 'missing' housing units or who live in enumerated housing units but were not detected by the enumerators had no chance for enumeration in the DOT Travel-to-Work Supplement. The person ratio estimation corrects for these deficiencies with respect to the count of persons 14+ in each SMSA. Likewise, the second stage ratio estimation was employed to reduce the effect of the these deficiencies, although some bias in AHS data still exists.

Table B. Estimates of the Number of Housing Units Missed by the 1975-76 Annual Housing Survey, for 21 SMSA's

	Estimated Conventional New Construction Units Missed by 1975 AHS-SMSA Survey	Estimated New Mobile Homes Missed by 1975 AHS-SMSA Survey
SMSA		
Atlanta Chicago Cincinnati Colorado Springs Columbus Hartford Kansas City Madison Miami Milwaukee New Orleans Newport News- Paterson- Philadelphia Portland Rochester San Antonio San Bernardino- San Diego San Francisco- Springfield-	10,100 34,400 2,300 1,800 3,700 4,700 4,700 1,500 17,100 4,000 3,500 1,400 2,100 18,800 5,100 2,800 2,900 3,100 8,900 7,200 1,500	5,200 4,800 1,700 1,900 1,800 300 1,200 500 5,300 1,000 1,800 800 0 2,200 4,400 700 400 10,300 13,400 2,300 300

Rounding errors With respect to errors associated with processing, the rounding of estimates introduces another source of error in the data, the severity of which depends on the statistic being measured. The effect of rounding is significant relative to the sampling error only for small percentages and medians derived from relatively large bases (e.g., median number of workers per household or median distance traveled to work).

This means that confidence intervals formed from the standard errors given may be distorted, and this should be taken into account when considering the results of the survey.

Sampling errors The particular sample used for this survey is one of a large number of possible samples of the same size that could have been selected using the same sample design. Even if the same schedules, instructions, and enumerators were used, estimates from each of the different samples would differ from each other. The deviation of a sample estimate from the average of all possible samples is defined as the sampling error. The standard error of a survey estimate attempts to provide a measure of this variation among the estimates from the possible samples and, thus, is a measure of the precision with which an estimate from a sample approximates the average result of all possible samples.

As calculated for this survey, the standard error also partially measures the variation in the estimates due to response and enumerator errors (nonsampling errors), but it does not measure, as such, any systematic biases in the data. Therefore, the accuracy of the estimate depends on both the sampling and nonsampling error measured by the standard error, biases, and some additional nonsampling errors not measured by the standard error.

The sample estimate and its estimated standard error enable the user to construct interval estimates in which the interval includes the average result of all possible samples with a known probability. For example, if all possible samples were selected, each of these surveyed under essentially the same general conditions, and an estimate and its estimated standard error were calculated from each sample, then:

- 1. Approximately 68 percent of the intervals from one standard error below the estimate to one standard error above the estimate would include the average result of all possible samples.
- 2. Approximately 90 percent of the intervals from 1.6 standard errors below the estimate to 1.6 standard errors above the estimate would include the average result of all possible samples.
- 3. Approximately 95 percent of the intervals from two standard errors below the estimate to two standard errors above the estimate would include the average result of all possible samples.

For very small estimates the lower limit of the confidence interval may be negative. In this case, a better approximation to the true interval estimate can be achieved by restricting the interval estimate to positive values, that is, by changing the lower limit of the interval estimate to zero.

The average result of all possible samples either is or is not contained in any particular computed interval. However, for a particular sample, one can say with specified confidence that the average result of all possible samples is included in the constructed interval.

The figures presented at the end of this appendix are approximations to the standard errors of various estimates for each SMSA. In order to derive standard errors that would be applicable to a wide variety of items and also could be prepared at a moderate cost, a number of approximations were required. As a result, the tables of standard errors provide an indication of the order of magnitude of the standard errors rather than precise standard errors for any specific item.

Tables presented at the end of this appendix contain standard errors applicable to estimates of housing characteristics and travel—to—work characteristics of persons 14 years and older who were employed at the time of the 1975—76 AHS—SMSA survey. Tables I and II present standard errors for estimated levels of housing characteristics, while tables III through VI present standard errors for estimated levels of travel—to—work characteristics. Tables VIII and IX present standard errors for estimated percentages for housing and travel—to—work characteristics, respectively. The figures given in tables VIII and IX cannot be used without applying the factors shown in table VII. Included in these tables are estimates of standard errors for estimates of zero and zero percent. These estimates of standard errors are considered as over—estimates of the true standard errors and should be used primarily for construction of confidence intervals for characteristics when an estimate of zero is obtained.

For ratios, 100 x/y, where x is not a subclass of y, the percentage tables for each SMSA underestimate the standard error of the ratio when there is little or no correlation between x and y. For this type of ratio, a better approximation of the standard error may be obtained by letting the standard error of the ratio be approximately equal to:

$$(100) \quad \frac{x}{x} \sqrt{\frac{\sigma_x}{x}} + \left(\frac{\sigma_y}{y}\right)^2$$

where: x = the numerator of the ratio

y = the denominator of the ratio

σ = the standard error of the numerator

 $\sigma_{\mathbf{y}}$ = the standard error of the denominator

Illustration of the Use of the Standard Error Tables

The results of the DOT Supplement indicate that in the Rochester SMSA 197,200 persons commuted to work in 1975. Interpolation in table IV of this appendix for the Rochester SMSA shows that the standard error of an estimate of this size is approximately 3,280. Consequently, the 68-percent confidence interval, as shown by these data, is from 193,920 to 200,480. Therefore, a conclusion that the average estimate, derived from all possible samples, of 1975 persons who commuted to work lies within a range computed in this way would be correct for roughly 68 percent of all possible samples. Similarly, we could conclude that the average estimate, derived from all possible samples, lies within the interval from 191,950 to 202,450 workers with 90-percent confidence and within the interval from 190,640 to 203,760 with 95-percent confidence.

Also, of the 197,200 workers who commuted 138,200 or 70.1 percent drove alone. Interpolation in table IX and application of the factor for Rochester in table VII show that the standard error of the 70.1 percent is approximately 1.08 x .860 = .9 percentage points. Consequently, the 68-percent confidence interval, as shown by these data, is from 69.2 to 71.0 percent; the 90-percent confidence interval is from 68.7 to 71.5 percent; and the 95-percent confidence interval is from 68.3 to 71.9 percent.

Standard Errors of Differences. The standard errors shown are not directly applicable to differences between two sample estimates. The standard error of a difference between estimates is approximately equal to the square root of the sum of the squares of the standard error of each estimate considered separately. This formula is quite accurate for the difference between estimates of the same characteristic in two different SMSA's or the difference between separate and uncorrelated characteristics in the same SMSA. However, if there is a high positive correlation between the two characteristics, the formula will overestimate the true standard error; whereas, if there is a high negative correlation, the formula will underestimate the true standard error.

Illustration of the Computation of the Standard Error of a Difference In 1975 there were 41,100 persons who carpooled to work in the Rochester SMSA. Thus, the apparent difference, as shown by these data, between those commuters who drove alone and those who carpooled in 1975 is 97,100. Table IV for Rochester shows the standard error of 138,200 is approximately 2,910 and the standard error of 41,100 is approximately 1,710. Therefore, the standard error of the estimated difference of 97,100 is about

$$3,380 = \sqrt{(2,910^2 + (1,710)^2)}.$$

Consequently, the 68-percent confidence interval for the 97,100 difference is from 93,720 to 100,480 workers. Therefore, a conclusion that the average estimate of this difference, derived from all possible samples, lies within a range computed in this way would be correct for roughly 68 percent of all possible samples. Similarly, the 90-percent confidence interval is from 91,690 to 102,510, and the 95-percent confidence interval is from 90,340 to 103,860. Thus, we can conclude with 95-percent confidence that the number of commuters who drove alone in 1975 is greater than the number who carpooled, since the 95-percent confidence interval does not include zero or negative values.

Standard Error of an Arithmetic Mean.—The standard error of an arithmetic mean can be approximated by the following formula:

$$\frac{\sigma}{x} = \sqrt{\frac{b}{y}} S^2$$

where y is the size of the base, and b is calculated by squaring the appropriate factor in table VII and multiplying by 101.66896 for housing data and 106.18734 for workers data. The variance, S², is given by

$$S^{2} = \sum_{i=1}^{i} p_{i} \vec{X}_{i}^{2} - \vec{X}^{2}$$

where \overline{x} is the mean of the distribution; c is the number of groups; i indicates a specific group, thus taking on values 1 through c; p, is the estimated proportion with the characteristic in group i; Z_i and Z_i are the lower and upper interval boundaries respectively,

for group i; and $\bar{x}_i = \frac{Z_{i-1} + Z_i}{2}$, which is assumed to be the most representative value for the characteristic for persons or families and unrelated individuals in group i. Group c is open-ended, i.e., no upper interval boundary exists. For this group an approximate average value is

$$\overline{x}_{c} = \frac{3}{2} Z_{c-1}$$

Standard Error of a Median. The sampling variability of an estimated median depends upon the form of the distribution as well as the size of its base. An approximate method for measuring the reliability of a median is to determine an interval about the estimated median, such that there is a stated degree of confidence that the median based on a complete census lies within the interval. The following procedure may be used to estimate the 68 percent confidence limits on sample data:

- (1) Determine, using the standard error tables and factors, the standard error of the estimate of 50 percent from the distribution.
- (2) Add to and subtract from 50 percent the standard error determined in step (1).
- (3) Using the distribution of the characteristic, calculate the confidence interval corresponding to the two points established in step (2).

A two-standard-error confidence interval may be determined by finding the values corresponding to 50 percent plus and minus twice the standard error determined in step (1).

Illustration of the Computation of a Confidence Interval for a Median. The following is the distribution of travel time from home to work for heads of owner-occupied housing units in Rochester as obtained from the DOT Supplement. These figures have been rounded to the nearest hundred.

Class Interval	Number of Persons	Relative Frequency (%)
Less than 15 min. 15 to 29 min. 30 to 44 min. 45 to 59 min. 1 hr. to 1 hr. 29 min. 1 hr. 30 min. or more	32,400 59,800 23,500 5,600 3,400 600	25.9 47.7 18.8 4.5 2.7
Total	125,300	•

The median, as computed from this distribution on a base of 125,300, is 22.1 minutes.

- (1) Using table IX and the factor for Rochester, the standard error of 50 percent on a base of 125,300 is about 1.2 percent.
- (2) A 95 percent confidence interval on a median is obtained by adding to and subtracting from 50 percent twice the standard error found in step (1). This yields percent limits 47.6 and 52.4.

(3) The median interval is 15 to 29 minutes. (For purposes of calculation the interval is taken to be 14.5 to 29.5 minutes.) It can be seen above that 25.9 percent of the persons fall in the interval below the median interval while 47.7 percent fall in the median interval itself. Thus, the lower limit on the estimate is found to be about

$$14.5 + (29.5 - 14.5) \left(\frac{47.6 - 25.9}{47.7} \right) = 21.3$$

Similarly, the upper limit may be found by linear interpolation to be about

$$14.5 + (29.5 - 14.5) \left(\frac{52.4 - 25.9}{47.7} \right) = 22.8$$

Thus, the 95 percent confidence interval on the estimated median is from 21.3 to 22.8 minutes.

Note that a median calculated directly from the data file will probably differ from that calculated using the distribution. In this case, form a new interval by adding to and subtracting from the directly calculated median the differences between the median and the endpoints of the confidence limits derived from the grouped data.

For example, suppose the median computed directly from the data file was 23 minutes. Then the 95-percent confidence interval would be from 22.2 [23 - (22.1 - 21.3)] to 23.7 [23 + (22.8 - 22.1)].

Illustration of the Computation of the Standard Error of an Arithmetic Mean.—In the previous example the distribution of travel time from home to work for heads of owner-occupied housing units in Rochester was given. The mean travel time, as calculated from this grouped data, is 24.8 minutes. The values of P_i and \overline{X}_i for each group are shown below:

Class Interval	P _i	<u> </u>
Less than 15 min. 15 to 29 min.	• 259	7.5
30 to 44 min.	•477 •188	22.5 37.5
45 to 59 min. 1 hr. to 1 hr. 29 min.	•045 •027	52.5
1 hr. 30 min. or more	.005	75.0 135.0

TABLE I. Standard Errors for Estimated Number of Housing Units in the 1975 Housing Inventory: Four large SNSA's

(68 thances out of 100)

2,000,000	1,500,000	1,000,000	500,000	400,000	250,000	100,000	50,000	25,000	10,000	5,000	2,500	1,000	700	500	200	0	Estimate	Size of			
- 1	;	1	4,260	3,950	3,280	2,170	1,560	1,110	710	500	350	220	190	160	100	50	SMSA	-			
1	;	;	:	:	:	1,260	1,060	800	530	380	270	170	140	120	80	30	City	In Central		Atlanta, Ga.	
1	;	;	!	3,910	3,530	2,480	1,810	1,300	830	590	. 420	260	220	. 190	120	70	City	Central	Not In		
16,440	15,030	12,890	9,530	8,590	6,880	4,400	3,130	2,210	1,400	990	700	440	370	310	200	200	SHSA				
1	:	7,360	7,870	7,430	6,300	4,240	3,060	2,180	1,390	980	700	440	370	310	200	190	City	In Central		Chicago, 111.	S
:	:	12,000	9,260	8,420	6,810	4,400	3, 130	2,200	1,410	1,000	710	450	370	320	200	200	City	Central	Not In		Standard Erro
1	11,300 ·	10,200	7,850	7,130	5,760	3,720	2,650	1,880	1,190	840	600	380	320	270	170	140	SMSA			Philad	ř
;	:	;	4,660	4,830	4,480	3,200	2,340	1,680	1,080	760	540	340	290	240	150	120		In Central		Philadelphia, Pa	
:	:	060,00	7,810	7,210	5,960	3,920	2,810	2,000	1,270	900	610	400	340	280	180	160	City	Central	Not In	N.J.	
1	!	6,020	5,870	5,500	4,620	3,080	2,220	1,580	1,010	710	500	320	270	230	140	100	SMSA			San F	-
:	;	;	;	2,900	3,270	2,540	1,900	1,370	880	630	440	280	240	200	130	80	Cities	In Central		San Francisco-Oakland, Cali	
:	;	!	5,510	5,410	4,790	3,320	2,420	1,730	1,100	780	550	350	290	250	160	120	Cities	Central	Not In	nd, Calif	

TABLE II. Standard Errors of Estimated Number of Housing Units in the 1975 Housing Inventory: Seventeen Smail SMSA's

(68 chances out of 100)

	000,000	250,000	100,000	50,000	25,000	10,000	5,000	2,500	1,000	700	50	2	10:3	Size of Estimate	VSIE
	3,640	3,990	3,070	2,290	1,650	1,060	750	530	340	280	246	<u>%</u>	0TT 0TT	Chiemac Onio- KyInd.	
	1	. 1	1990	a k o	710	26.	350	230	160	130	110	ક	55 GS	Colo.	
	ı ı	2,790	2,440	1,850	1,360	<i>87</i> 0	620	otto	280	230	8	130	88	Ohio	
	E (ŧ	1,740	1,400	1,050	680	490	350	22	160	160	100	70 70	ford, Jonn.	
	3,520	3,890	3,000	2,230	1,620	1,040	740	520	330	280	230	150	0TT 0TT	Oity, No:- Kans.	
	1 1	1	690	870	710	760	350	250	160	130	OTT	8	ઝઝ	Wis.	Sta
	3.230	4,700	3,590	2,670	1,930	1,240	86	620	390	330	290	180	160 160		Standard Brror
	, C	3,920	3,050	2,270	1,650	1,060	750	530	340	280	240	051	or Tro	Wis.	TOT
	2,090	3,250	2,170	2,050	1,490	%	680	490	310	88	22	T.	<u>8</u> 8	1.0	Ē
	l: 1	1	8	950	760	510	376	86	170	υt	120	8	75 75	News, Va.	
	, A		2,740	2,060	1,500	%	690	490	310	260	220	υtτ	ь 15 8	Clifton- Passalo, N.J.	Pataram
). }	3,83	2,870	2,130	1,540	980	700	500	OTE	86	220	140	100 100	Oreg.— Wash.	Port I and
	1 1	2,660	2,320	1,770	1,290	830	590	420	270	23	190	120	870	N.T.	Portland Rochester
	, ,	3,090	2,420	1,800	1,310	о т	8	0 2 4	270	220	190	120	8 %	Antonio Tex	S .
	2,990	3 4,060	3,080	2, 290	1,650	1,060	750	530	340	280	240	150	110 110 110	L 3 3 6	Sam
	3,210	4,630	3,530	2,630	1,900	1,220	86	OT9	390	320	270	170	150 150	58	San
186	1.00 I	ing the second s	1,590	1,280	950	83	M _t O	320	8	170	胺	8	88	Diego Chicopee- Calif Holyoke, Mass	Springfield

TABLE III. Standard Errors of Estimated Numbers of Workers in the Four Large SNSA's, the Central Cities, and the Balance of the SNSA's

(68 chances out of 100)

		Atlanta, Ga.	Not In		Chicago, 111.	No.+	Phila	7	aN.J.	San F	San Francisco-Oakland, Calif.
Size of		Central	Central		Central	Central		Central	Not In		
Detinote of	220			CMCA	2	711	CHCA	City	2011	3	2
De Characte	5	3:3	3.5	51.00	3.50	10.00	00000	Care	CTC		OFTION
0	50	30	70	220	200	230	150	110	170		110
25	50	33	70	220	200	230	150	110	170		110
50	50	40	70	220	200	230	150	110	170		110
100	70	60	80	220	200	230	150	110	170		110
200	100	80	120	220	200	230	170	150	180		150
500	160	120	190	330	310	340	270	230	290		230
700	190	150	220	390	370	400	320	270	340		270
1,000	230	170	260	470	440	480	380	330	• 10		330
2,500	360	280	410	750	700	760	600	510	650		510
5,000	510	39.0	590	1,050	990	1,070	850	730	920		730
10,000	720	550	820	1,490	1,400	1,510	1,200	1,020	1,300	ŗ	030
25,000	1,130	840	1,290	2,350	2,210	2,380	1,900	1,610	2,050	<u>.</u>	520
50,000	1,580	1,140	1,800	3, 320	3,100	3,350	2,680	2,260	2,880	2,2	80
75,000	1,910	1,330	2,170	4,060	3,780	4,080	3,270	2,740	3,510	2,7	8
100,000	2, 180	1,460	2,460	4,670	4,340	4,690	3,760	3, 130	4,030	3,1	90
150,000	2,610	1,570	2,910	5,700	5,260	5,700	4,570	3,760	4,870	3,870	70
250,000	3,200	1,310	3,460	7,280	6,630	7,220	5,820	4,660	6,140	4,8	80
500,000	3,830	!	3,660	10,030	8,810	9,700	7,920	5,820	8,110	6,5	8
1,000,000	2,550	;	;	13,410	10,640	12,170	10,260	5,450	9,650	7,9	20
1,400,000	;	:	;	15,090	10,570	12,740	11,170	!	9,360	7,9	6
2,000,000	;	;	;	16,550	7.,690	11,630	11,390	:	5,750	6	6,180
2,250,000	!	!	:	16,850	4,420	10,340	11,100	: 	!	<u>.</u>	4,300
2,500,000	;	:	1	17,000	;	8,280	10,570	:	1		:
3,500,000	;	;		15,950	:	:	3,880	:	:		!
4,500,000	;	;	!	11,580	:	;	1	:	¦		!
5,000,000	ţ	1	1	6,470	1	:	!	;	;		:

	•
(68 chances out of 104)	

TABLE 17. Standard Errors of Estimated Numbers of Workers: Seventeen Small SISA's

, .	-	٠		-	•														٠.								_	'			
., 000		900,000	800 000	600,000	500,000	400,000	300,000	250,000	200,000	150,000	100,000	75,000	2000	22,000	10,000	2,000	, 300	3	3 3	100	2 5	3 5	3 4	5 0			Estimate	ize of			
i						-																			120		2	Ohlo-Ky			
;		1	! !	;	;	:	:	:	;	1,040	1,240	1,200	1,080	830	\$50	397	280	100	150	130		. 2	: -		30		Col .	Springs	•		
:		: :	.,0,0	170	1 10	3.720	3,710	3,590	3,380	3,070	7,620	2,310	1,920	1,390	890	630	450	280	240	200	130	90		90	90		0110	-			
;		: :			60	2.210	2.550	2,690	2,630	2,460	2,140	1,910	1,600	1,160	750	530	380	240	200	170	110		3	8	60		Comm				
;	4,470	0,920	9,140		# 120	5 250	4.940	4,680	3,960	3,870	3, 260	2,860	2,370	1,700	1,080	770	540	340	290	240	150	120	120	120	120		Kans	Kensas			
;	:	;	;	1	1	;	;	;	ē	1,140	1,180	1,120	990	740	190	350	250	3	30	Ē	70	So	å	20	20		The state of			١	2
4,030	5,0/0	5,720	6,270	440	200	A 9	5 540	5.210	2.790	4, 260	3,560	3.120	2.580	1,850	1.170	830	590	370	310	265	170	071	140	140	ī	1.				7.00	andard E
2,790	4,150	4,950	5,660	5,090			7	. 830	4.450	3,960	3.320	2.910	2.410	1,720	1.100	730	550	350	29:1	250	160	120	120	120	120		Mi andree'			101	,
;	;	1,160	3, 890	1.000			4 270	4 080	3.810	3.430	7.900	2.560	2.120	1.530	970	690	490	310	260	220	140	100	700	10:	3	50.	Orleans,	No.			
:	:	:	:	:	:		}			_	1.240															Mem, va.	Membort				
2,150	3, 790	4,670	5,470	5,530	3,500	2,020	7,750	730	360	1,000	1.260	2,50	170	700	1.080	770	540	340	290	240	150	120	120	120	120	1.3.	Passaic,	Clifton-	Paterson-		
•	:	1,730	4,090	4,470	4,500	4,540	1,1,0	100	1 000		970	2.620	2.170	560	20.5	710	500	320	260	220	-	100	100	100	100	wasn.	Oreg	Port and,			
:	:	:									2.580															N.T.	Rochester,	'			
1	;	;	2,150	3,080	3,460	3,330	4,750	1,100	4 3 5 0 0	2 0 0 0	2 5 5 6	2 240	170	1.350	200	2	÷.	270	230	190	120	8	3	8	3	Техаз	Anton lo.	San			
1	;	3, 190	4,680	4,940	4,940	4,690				* 5.50	4 5	2 750	3 3 3 6 6	640	2	740	520	330	280	230	150	===	110	=	=	Callf.	Onterio,	Riverside-	San San		
1,166	5.160	5,790	6, 320	6,280	6,040	5,570	21.240		41.200	3,300	04140	2, 390	1000	1,100				170	110	270	170	100	140	10	.	Calif.	Dicgo.	Son			
	:	:	:		300	1,750	1,940	2,000	1,940	1,/30		1,320	. 9/0	620	1	920	200	376			3 3	\$ 6	•	•	•	Comm.	Ma95	Holyoke,	Springfield-		

į			_	-	_					_	_	_	_	****	_	_	_	_		_		_	_	_			
	300,000	3	20,000	300,000	250,000	200,000	150,000	100,000	75,000	50,000	25,000	10,000	5,000	2,500	1,000	196	S00	200	2	S	25	•	23 23 25 20 20 20 20 20 20 20 20 20 20 20 20 20	10 9212	}		
	••		:	1,260	2,410	2, 870	2,980	2,780	2,540	2, 180	1,610	1,050	750	530	340	280	240	150	110	1	110	===	100.	Un10-Ky	Cincipnati.		
			1	:	;	:	:	480	910	870	730	500	360	260	170	5	120	70	50	å	30	8	cete.	Springs.	Colorado		
	:		:	2,360	2,660	Z. 750	Z, 670	Z, 390	2, 160	1,830	1,340	860	620	440	pre	230	200	120	3	80	8	8	OLINO	Calumbus,			
	:		:	;	;	;	ì	130	1,030	1,180	1,020	710	510	370	236	200	170	=	70	3	3	8	LONG.	Hartford,			
	:		:	2,410	2,980	3, 220	3, 190	2,900	2,630	2,240	1,650	1,060	760	540	34	290	240	150	120	120	120	120	XENS#S	C1ty, 20	Kansa s		
	:		:	:	1		;					_		240			_		_		_	_	#15C.	Madison			s
	:		;	;	2,280	2,920	3, 110	2,940	2,700	2,330	1,730	1,120	806	570	¥	38	250	160	130	1.30	130	130	718.	H) sec			Standard Error
	380	9,100	3	3.780	3,860	3,780	3,540	3,090	2,750	2,310	1,680	. oeo	770	540	-	296	249	150	120	120	120	120	WASC.	Mi Iwaukee,			Error
		20.00	-	2,890	3,130	3, 180	3,050	2,710	2,430	2,060	1,500	970	690	190	4 6	260	220	140	2	2	3	3		Or leans,	New		
	:		:	:	:	:	980	1,200	1, 180	1,060	810	540	390	270	77	150	120	*	8	à	30	5	News, Va.	Newport			
	:		;	:	:	970	2,100	2,330	2,240	1,990	1,510	990	710 ·	510	100	270	230	-	3	100	<u>.</u>	9	N.J.	Passaic,	Clifton-	Paterton.	
	;		:	:	1,760	2,500	2,740	2,630	2,430	2,100	1,560	1,020	720	818	3 3	270	210	150	110	110	=======================================	=	Wash.	Dreg	Portland		
	:	;	!	:	;	;	1,570	1,840	1,780	1,600	1,220	800	580	410	260	220	3	120	3	70	70	70	N.Y.	Rochester,			
	;	2,190	3	2.800	2,880	2,830	2,660	2,330	2,080	1,750	1,270	820	580	-	260	220	3	126	3	70	70	70	Техаз	Antonio,	Sen		
	ł	:		:	:	1,720	2,380	2,460	2,330	2,040	1,540	1.010	720		7/4	170	3 15	5	-	1	-	1	Callf.	Onterio.	Riverside-	San	
	2,670	4,030		a a	4.480	4, 320	000	3,460	3,080	2,580	1,870	1.200	850	500	320	•	110	170	5 5	5 5	5 8	Ē	Callf.	Di ego.	S		
	:	:	- 1	}				_						300			_	_	_				Mass. Conn.	ilol yoke.	Chicones-	?	

(68 chances out of 100)

TABLE VI. Standard Errors of Estimated Numbers of Norkers: Seventeen Small SHSA's: Bainnee

			-
\$60,000	ž .	Size of	
7.500 7.500		Clucionati Ohlo-Ky	
400 400 400 400 400 400 400 860 860 860 860	*	Colorado Springs,	
900 2000 2000 2000 2000 2000 2000 2000	8 8	Co hambus,	
500 100 100 100 100 100 100 100 100 100	8 8	Hartford,	
120 120 120 120 120 240 240 240 240 240 240 240 240 240 2	120	Kansas City, Mo	
1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1		Madison	
140 140 140 140 140 140 140 140 140 140	140	M.	Standard Error
139 119 119 119 119 119 119 119 119 119	130	Hi i waukee,	Error
100 100 100 100 100 100 100 100 100 100		New Orleans,	
111111111111111111111111111111111111111	News, Va.	Newport	
120 120 150 150 150 240 240 340 340 4,690 2,360 3,240 3,240 3,240 4,850 4,850 4,850 4,850 1,890	N.J.	Paterson- Clifton- Passaic,	
100 100 140 140 220 220 244 320 710 710 710 710 710 710 710 710 710 71	Wash.	Portland, Oreg	
1 1 30 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0	N.Y.	Rochester,	
1,450 1,450 1,450	Tex.	San Antonio,	
110 110 110 120 120 140 140 140 140 140 140 140 140 140 14	Callf.	San Bernadino- Rivorside- Ontario,	
1,140 1,140 1,140 1,140 1,140 1,140 1,140 1,140 1,140 1,140 1,140 1,140 1,140	Callie 140		
50 70 90 150 180 330 460 650 650 650 1,430 1,430 1,430	1	Springfjeld- Chicopes- Iblyoke,	

(68 chances out of 100)

TABLE VII. Factors to be Applied to Tables VIII and IX

	Tra	vel-to-Wor	k Data ¹		Housing Da	ta²
		Central			Central	
SMSA	SMSA-	City	Balance	SMSA	City	Balance
Atlanta	.702	.538	.805	.701	F 7F	026
Chicago	1.448	1.362	1.467)	.535	.826
Philadelphia	1.171	.998	1.265	1.392	1.380	1.400
San Francisco	1.000	.889	and the second s	1.182	1.073	1.263
Cincinnati	1.054		1.061	1.000	.883	1.100
		1.032	1.034	1.062		
Colorado Springs		.512	.579	.504		
Columbus	.867	.854	.900	.877		
Hartford	.733	.723	.703	.690		
Kansas City	1.057	1.047	1.056	1.038		
Madison	.484	.479	.501	.499		
Miami	1.145	1.106	1.165	1.238		
Milwaukee	1.071	1.058	1.088	1.058		
New Orleans	.951	.950	.961	.964		
Newport News	.519	.537	.479	527		
Paterson	1.054	.989	1.055	.965		
Portland	.972	1.003	.974	.984		
Rochester	.860	.801	.874	.835		
San Antonio	.842	.802	.851	.842		
San Bernadino	1.018	1.001	1.038	1.060		==
San Diego	1.150	1.174				
Springfield	.613	.580	1.132	1.218		
Obtrustrein	.013	.300	.645	.629		

¹ Factors given below should be applied only to standard errors in table IX.

Factors given below should be applied only to standard errors in table VIII.
--Not applicable.

TABLE VIII. Standard Errors for Estimated Percentages of Housing Units in the 1975 Housing Inventory 1

Base of			Estima	ted Percentag	e	·
Percentage	0 or 100	1 or 99	5 or 95	10 or 90	25 or 75	50
200 500 700 1,000 2,500 5,000 10,000 25,000 100,000 250,000 500,000 700,000 1,000,000	33.7 16.9 12.7 9.2 3.9 2.0 1.0 0.4 0.2 0.10 0.04 0.02 0.01	33.7 16.9 12.7 9.2 3.9 2.0 1.0 0.6 0.4 0.3 0.2 0.14 0.12 0.10	33.7 16.9 12.7 9.2 4.4 3.1 2.2 1.4 1.0 0.7 0.4 0.3 0.3 0.3	33.7 16.9 12.7 9.6 6.0 4.3 3.0 1.9 1.4 1.0 0.6 0.4 0.4	33.7 19.5 16.5 13.8 8.7 6.2 4.4 2.8 2.0 1.4 0.9 0.6 0.5	35.6 22.5 19.1 15.9 10.1 7.1 5.0 3.2 2.3 1.6 1.0 0.7 0.6 0.5

¹ The standard errors in this table cannot be used without applying the factors given in table VII.

Base of			Estimated	ated Percentag	ge .	
Percentage	0 or 100	1 or 99	5 or 95	10 or 90	25 or 75	50
100	51.5	51.5	51.5	51.5	۲1 x	E1 E
200	34.7	34.7	34.7	34.7	7 7	36 /
500	17.5	17.5	17.5	17.5	20 0	37.0
700	13.2	13.2	13.2		16.0	10.0
1,000	9.6	9.6	9 6	0 6	1 4 1	14.5
2,500	4.1	4.1	4.5	٠, ٠, ٠, ٠, ٠, ٠, ٠, ٠, ٠, ٠, ٠, ٠, ٠, ٠		10.5
5,000	2.1	2.1	3.2	4.4	, d	7 7
10,000	1.1	1.1	2.2	3 .	љ (Π `• 3 ὑ
25,000	0.4	0.6	1.4	2.0	2 2	, r
50,000	0.2	0.5	1.0	1.4	2.0	<i>ی</i> ر
75,000	0.14	0.4	0.8	<u> </u>	1.6	- : 0
100,000	0.11	0.3	0.7	1.0	1.4	7 .
150,000	0.07	0.3	0.6	0 .8	1.2	7 .
250,000	0.04	0.2	0.4	0.6	0.9	1.0
500,000	0.02	0.15	0.3	0.4	0.6	0.7
800,000	0.01	0.11	0.3	0.3	0.5	0.6
1,000,000	0.01	0.10	0.2	0.3	0.4	о с
1,600,000	0.01	0.08	. 0.2	0.2	0.4) \ \
2,000,000	0.01	0.07	0.2	0.2	0.3	- t
2,250,000	0.01	0.07	0.15	0.2	0.3	0.3

¹ The standard errors in this table cannot be used without applying the factors given in table VII.